| DISTRICT | Aurora |
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| | |
| | |
| DIST_NO | 0410 66060510 |
| COUNTY If different from written on document | Mineral |
| TITLE If not obvious | Study, East Humboldt Vein, Aurora, Nevada, 1981 Ore blocks |
| AUTHOR | M.W. Cassiday |
| DATE OF DOC(S) MULTI_DIST Y / N? | 1982 52 pgs |
| Additional Dist_Nos: | |
| QUAD_NAME | Aurora 7.5' |
| P_M_C_NAME (mine, claim & company names) | East Humboldt Vein; Miller-Kappes Co.; Vinnel Corp. |
| COMMODITY If not obvious | gold; silve |
| NOTES | Feasibility study; property report; assays; reserves; |
| | NOTE do not scon last 11p - copyrighted |
| Keep docs at about 250 pages in (for every 1 oversized page (>11 the amount of pages by ~25) | Ix17) with text reduce DB: Initials Date DB: Initials Date SCANNED: T.M. 4/4/09 WT 12/21/09 |
| Revised: 1/22/08 | Initials Date |

PRELIMINARY FEASIBILITY STUDY

EAST HUMBOLDT VEIN

AURORA, NEVADA

- 1981 ORE BLOCKS -

Miller-Kappes Company

P. O. Box 13687, Reno, Nevada 89507 702-356-7107
1845 Glendale Avenue, Sparks, Nevada 89431
20 January, 1982

PRELIMINARY FEASIBILITY STUDY

EAST HUMBOLDT VEIN, AURORA - 1981 ORE BLOCKS -

INTRODUCTION AND GEOLOGIC SETTING

This report presents a very brief evaluation of the economics of an open pit mine and heap leach on a near-surface ore block drilled out in the fall of 1981 (hereafter called ore block 1981 A), on the east Humboldt Vein System, Aurora, Nevada.

Figure 1 presents a large scale map of the property showing the proposed open pit and the proposed heap leach site. In the initial 1982 mining program, a total of 236,000 tons of material - 56,000 tons of ore and 180,000 tons of waste - will be mined. Waste will be dumped near the mine and ore will be hauled approximately one mile to a stockpile at the heap leach site. Stockpiled ore will be crushed to 1/2-inch in a closed-circuit, two-stage crusher, stacked on plastic pads 15 feet deep, and heap leached. Based on extensive test work (Appendix A), estimated recoveries will be 65 percent of fire-assayable gold, in approximately two months of leaching. Some minor additional recovery might be possible after the heap sits dormant during the winter months.

The ore zone consists of two parallel, continuous quartz veins that dip about 70 degrees to the west. An area designated "Ore Block A" on the accompanying map contains proven ore which has been extensively drilled and sampled. A second area of possible ore, which will be excavated as part of the access ramp into the proposed pit, has been designated "Ore Block B". Figure 2 is a plan view on a scale of 1" = 20', and Figures 3 through 9 are ore reserve cross-sections, as prepared by Bruce W. Miller assisted by Joanne Seegelken and Andy Glatiotis. Based on an arbitrary pit depth of 100 feet, the overall ore tonnage which he calculates in ore block 1981 A, is 41,224 tons, with an average grade of 0.116 ounce gold per ton. Stanley Reamsbottom, a geologist in Vancouver, B.C., independently calculated a slightly larger pit to contain 49,503 tons of ore, assaying 0.105 ounces gold.

per ton. The two pit designs contain within 8 percent of the same amount of gold, and in this study, ore block 1981 A will be assumed to contain 41,000 tons of open pit mineable ore, with a head grade of 0.116 ounces gold per ton.

Because of the pit design, it seems to be overly conservative to base the financial analysis on only the proven reserves mentioned above. The West vein, as it extends north from the pit, is exposed on the surface as a strong, continuous quartz outcrop for several hundred feet. The final pit design calls for construction of a pit access road from the north, which, during construction, will mine out over 15,000 tons of this West vein material lying north of the proven ore reserves. Surface and limited underground sampling indicate that this 15,000 ton zone (designated ore block 1981 B) is lower grade than the target ore block. However, it seems reasonably safe to impute an ore grade of, at least, 0.05 ounces gold per ton, over the 15 foot projected width (haulage road cross-sections shown in Figures 10 through 14) that was used to calculate the tonnage. Ore reserves in this area are still classified as potential by the project geologists. Thus, the ore may not materialize. To allow for this probability, after calculation of the feasibility analysis, a short analysis was run, assuming only ore block A is present. That analysis shows that the overall project will still show a slight operating profit, hence, it appears justified to base the mine design and economic projections on the assumption that this ore zone is there.

ADDITIONAL MINING

This feasibility study is based on two ore blocks, containing 41,000 tons of proven ore, and another 15,000 tons of possible ore. Geologic projections previously made for the property (see S. Reamsbottom report dated April 3, 1981), indicate that there are other potential ore reserves within 100 feet of the surface, which may be open-pittable under economic conditions similar to those present here. At-depth reserves, greater than 100 feet deep and accessible by underground mining, also represent considerable tonnage, which is currently classified as probable, and has not been considered as part of this feasibility study.

It is expected that mining will commence in 1982 on the limited pit proposed here, with all mining and most leach operations completed by the end of 1982. Additional geologic evaluation of other reserves should lead to a continuation of mining in subsequent years.

MINING

DESCRIPTION OF OPERATION

Access to the orebody for mining will require the construction of two access roads onto the orebody, as shown in orange in Figure 1. The natural outcrop of the ore zone ranges in elevation from 7380 feet to about 7460 feet. (In Figures 2 through 14, add 6200 feet to the elevation shown to get the approximate true elevation corresponding to Figure 1.) The upper access road cuts the ore zone at elevation 7400, and the lower access road cuts the ore zone at elevation 7300. Minor offsets from these roads can be used to access the orebody on intermediate 20 foot levels. The upper road will be used for haul truck access for ore only. Detailed pit designs have not yet been developed, but it is thought that all waste from the upper levels of the ore zone can be moved by bulldozer for access from the 7300 foot road level. The two roads can be placed along existing old roads, or along contours where little or no cut and fill work is needed, except to level the road as it traverses the hillsides.

Road construction will require an estimated 10 days of D-8 bulldozer work at \$100.00 per hour, and 5 days of road-grader work at \$40.00 per hour, for a total cost of about \$10,000.00.

Mining Methods. Larie Richardson, of Vinnel Corporation, has suggested that it would probably be a mistake to perform significant bulldozer operations within such a small pit. He feels that mixing of ore and waste is inevitable in that case, and suggests that nearly all waste be lightly blasted and hauled.

Since the vein material is readily distinguishable, waste, in the wide upper third of the pit (above elevation 7360), could probably be ripped and pushed, but all the ore will be drilled and blasted. Ore is uniform enough to permit mining on 20 foot vertical benches, however, equipment selection may dictate a shorter bench height. Pit sideslopes, and the south end wall slope, will be 55 degrees, resulting in a 14 foot horizontal run for each 20 foot high bench to provide a wide cleanup road. Alternate benches, for instance, benches at levels 1240, 1200 and 1160 will be moved to within 5 feet of the wall, and the intermediate benches, 1220, 1180, etc., will be left as horizontal roadways. If benches shallower than 20 feet are used, the same approximate procedure will be employed, leaving 20 foot wide cleanup roads approximately every 40 vertical feet.

Three contractors have provided preliminary estimates of cost, with all costs generally in the \$2.00 per ton range for either ore or waste, based on a haul of up to two miles for both ore and waste.

Approximately 20,000 tons of waste, in the area of the 7300 foot level pit access road, and the outer edges of the pit, will be ripped and bulldozed to the dump site. Costs for this waste should be less than \$1.00 per ton. The bulk of the remainder of the waste will be truck-hauled 1000 feet to a waste dump located immediately adjacent to the 7300 foot level access road on the north pit edge. Cost to move this 166,000 tons of waste should result in a cost of \$1.70 - \$0.30 per ton lower than the longer haul on which the estimated bids were based.

Approximately 15 percent of the total ore tonnage exists as a rib of quartz sticking above the present ground surface. This rib is, in many places, a jumble of huge boulders. Mining this ore will require drilling and blasting of many boulders, or alternatively, "mud capping" them with shaped charges. This will add an estimated \$2.00 per ton to this ore, or \$0.30 per ton to overall ore costs.

The following cost estimates for mining and hauling of ore and waste are used in the feasibility.

| Ore hauled to stockpile at heap leach (1 mile haul) | \$2.30 per ton |
|---|----------------|
| Waste pushed or hauled to north end dump | \$1.60 per ton |
| Waste hauled to a remote dump | \$2.00 per ton |

CRUSHING

All ore will be crushed in a closed-circuit system to one half inch. Contractors estimates for crushing costs, including stacking via a long segmented conveyor ending in a mobile stacker, have been in the range of \$3.00 to \$3.50 per ton. One California crusher operator suggested mobilization for 30,000 tons of ore would result in a total cost of \$9.00 per ton. However, both Vinnell, who has worked on similar projects in the area, and Hunewill Construction Company at Wellington, who have a crusher currently idle, have quoted the lower figure for crushing and stacking.

A cost of \$3.50 per ton is used.

CONTRACTOR'S MOBILIZATION

Contractor's mobilization costs will range from approximately \$10,000.00 for Hunewill from Wellington (estimated, not confirmed), to \$75,000.00 for Vinnell. The higher figure is used in the study since Hunewill is not familiar with gold mining and will, no doubt, incur some startup problems.

LEACHING AND GOLD RECOVERY

Leach tests indicate the ore is clean and free of cyanicides. Leaching should be straightforward and predictable. Activated carbon will be used to recover gold from pregnant leach solutions. The carbon will then be stripped and the gold plated onto steel wool cathodes which will be smelted to produce a dore' bar.

Cost estimates to conduct the leach on the total 56,000 tons of ore are as follows:

| Pad and Pad Preparation | \$ 0.45 | per | ton |
|---|------------|-----|-----|
| Ponds | \$ 0.05 | per | ton |
| Fences and Roads | \$ 0.20 | per | ton |
| Chemicals, Lime and Cyanide | \$ 0.80 | per | ton |
| Carbon | \$ 0.25 | per | ton |
| Pipes and Miscellaneous Supplies | \$ 0.30 | per | ton |
| Labor, 3 Men @ \$2,500.00 per Man per Month, 4 Months duration | \$ 0.55 | per | ton |
| Power Generated On-Site or Purchased | | | |
| @ \$0.10 per KW Hour | \$ 0.20 | per | ton |
| Water pumped from Humboldt Mine workings via 3000 feet of 2-inch PVC pipeline, including electric pump and generator, | | | |
| Total - \$10,000.00 for equipment plus operating costs | \$ 0.20 | per | ton |
| Total Leaching Costs, including all services and supplies | \$ 3.00 | per | ton |

TOTAL OPERATING COST

Total operating cost to mine and treat the 56,000 tons of ore and 180,000 tons of waste are as follows:

| Haul Roads | \$ 10,000.00 |
|--|---------------|
| Mining: Ore, 56,000 Tons @ \$2.30 per ton; Waste, at Mine Dump, | 129,000.00 |
| 180,000 Tons @ \$1.60 per ton; Waste, Remote Dump, | 288,000.00 |
| -0- Tons @ \$2.00 per ton | -0- |
| Crushing and Stacking: 56,000 Tons @ \$3.50 | |
| per ton; | 196,000.00 |
| Mobilization: | 75,000.00 |
| Leaching: 56,000 Tons @ \$3.00 per ton; | 168,000.00 |
| Miscellaneous Supplies; | 5,000.00 |
| Professional Time and Project Management; | 60,000.00 |
| TOTAL | \$ 931,000.00 |

CAPITAL COST

Since mining, crushing and stacking are contracted for, capital items for heap leach are included in operating costs, and a camp already exists at the property, there are very few actual capital items. Among the capital items necessary, and not considered elsewhere, are two pick up trucks, better communication facilities, additional office and safety equipment, and an on-site assay laboratory and the recovery plant.

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These items are all portable or re-saleable at the completion of the project. The items include:

| On-site Assay Laboratory | \$ 40,000.00 |
|------------------------------------|------------------|
| Recovery Plant | 70,000.00 |
| Pickup Trucks | 20,000.00 |
| Safety Equipment and Miscellaneous | 20,000.00 |
| Communications Facilities | 10,000.00 |
| | |
| Total Capital Items | \$ 160,000.00 |

Since all expenses will be made essentially before production begins, the entire capital plus operating expenses are considered, for financial purposes, to be the capital requirement.

PROFITABILITY

The following table presents a summary of the feasibility economics for mining ore blocks 1981 A and B on the East Humboldt Vein, Aurora, Nevada. Total project cost will be \$1,091,000.00, with a production revenue of \$1,432,600.00, leaving a \$341,000.00 net profit before taxes.

CAPITAL COST

| operating Expenses | |
|-------------------------|---------------|
| Operating Expenses | 931,000.00 |
| Capital Equipment Items | \$ 160,000.00 |

PRODUCTION REVENUE (@ 65 Percent Recovery and \$400.00 U.S. per ounce gold)

| 41,000 Tons @ .116 oz per ton | \$ 1,237,000.00 |
|---|-----------------|
| 15,000 Tons @ .05 oz per ton | 195,000.00 |
| Total | \$ 1,432,000.00 |
| Cash Flow through end of 1982 | \$ 109,000.00 |
| Net Return over Cost: Cash Flow through Project Completion in August, 1983 ⁽¹⁾ | \$ 341,000.00 |

NOTE: All costs and revenues in this report are shown in U.S. dollars at a gold price of U.S. \$400.00 per ounce. All costs and revenues are shown before taxes.

(1) Additional cash flow would accrue from residual value of capital equipment.

OPERATING CAPITAL REQUIREMENTS AND CASH FLOWS

The project will begin in March, 1982 with completion of mining and heap stacking by September of 1982. Leaching will commence in August, 1982 and it is anticipated that two months of 24-hours per day leaching, and two months of 8-hours per day leaching will be possible before winter temporarily halts the operation. Recoveries through the end of 1982 are expected to be a minimum of \$1,200,000.00, with the remaining \$31,000.00 being recovered in three months of leaching in early 1983.

SENSITIVITY AND STATISTICS

A statistical study was run to determine the validity of the ore grade determination and is included as Appendix B. Using the worst of 5 cases studied, a lognormal distribution with two high assay values thrown out and zero values excluded, the lower limit of ore grade at the 90 percent confidence level is 0.098 ounce gold per ton.

A sensitivity analysis is included at the end of Appendix B presenting several possible scenarios. One of these uses a gold price of \$360.00 per ounce and the minimum ore grade from the statistical study. This analysis shows the venture to still have higher returns than costs.

An alternate sensitivity scenario analyzes the effect on costs if the same total ounces of gold are recovered as originally expected, at a price of \$360.00 per ounce and if due to dilution, an additional 20 percent of "waste" is diverted to the heap to be crushed and leached. The venture shows positive returns here also.

SUMMARY

Since the feasibility study presented here shows a net, before tax, profit of \$341,000.00; and since capital improvements, such as the leach plant, water system, improved communications, pickup trucks and operating expenses all represent valuable assets to a continued mining operation

which is anticipated, the project appears to be justified. We recommend that negotiations for acquisition of the heap leach sites amd right-of-access for haul roads be begun immediately, and that funds be committed for finalizing the feasibility study and beginning the necessary permit acquisition. Construction of the carbon recovery circuit would be started in March, with the goal of starting preliminary production work on the property in late spring, as soon as weather and conditions are favorable.

Submitted by,

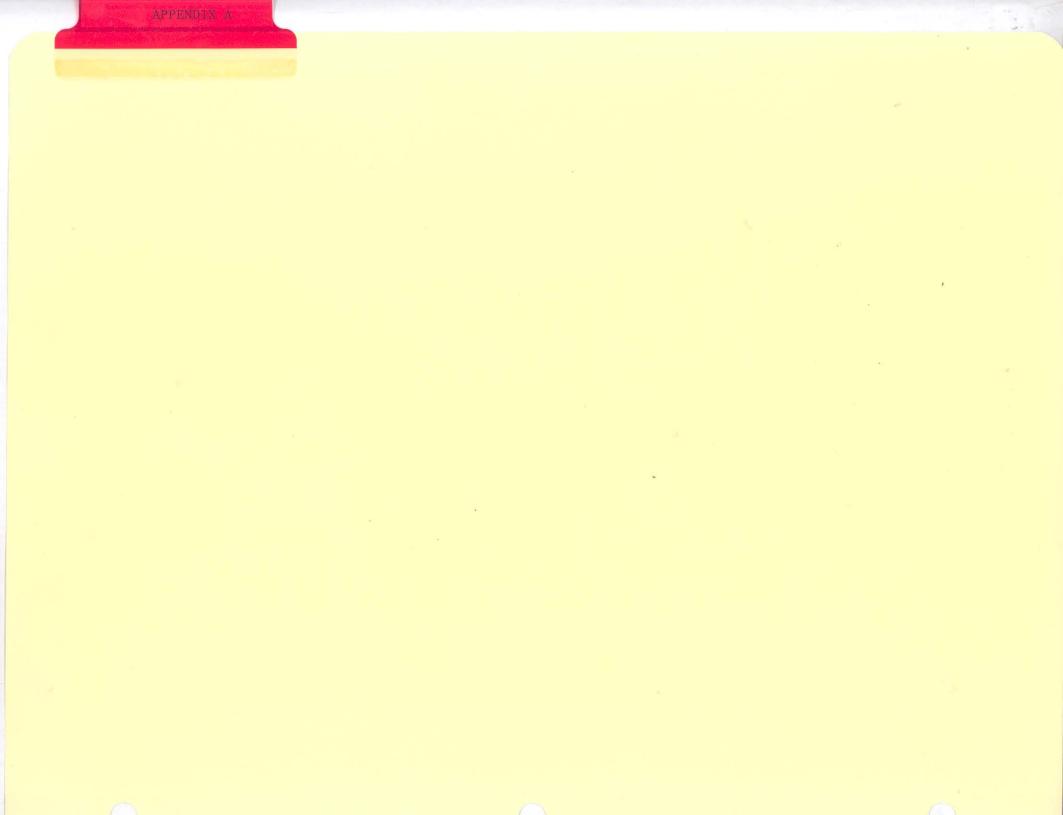
Michael W. Cassiday MILLER-KAPPES COMPANY

MWC/df

Attachments: Appendix A: Bucket Leach Tests on Aurora Samples

Appendix B: Statistical Analysis of Sample Data

and Sensitivity Study



APPENDIX A

BUCKET LEACH TESTS ON AURORA SAMPLES

This report summarizes the results of a testing program on nine samples of ore taken from the Aurora, Nevada property.

Figure 1 presents sample descriptions and head assays. A summary of the bucket leach test results is presented in Figure 2. Of the nine samples used for bucket leach tests, one sample (1759B) contained only 0.010 ounces gold per ton, with the remaining seven samples averaging 0.161 ounces gold per ton.

Recoveries, overall, were good, with 64 percent of contained gold, on average, being recovered. Recoveries from tests on ore crushed to 2-inches and 1-1/2-inches (7 tests), averaged 59 percent of contained gold. Recoveries from tests on ore crushed to 5/8-inch (6 tests), averaged 72 percent of contained gold.

SAMPLE PREPARATION

Sample 1701 A, B and D. The as-received samples, each weighing approximately 20 Kg, were set up as bucket leach tests. These samples are discussed later in the Test Histories section.

Sample 1701 C. The as-received sample weighed approximately 10 Kg and was treated as follows:

- 1) Crush entire sample through jaw crusher to 5/8-inch.
- 2) Split out 3 Kg of 5/8-inch material, using a Jones splitter, and set up a leach test.
- 3) Crush remaining 5/8-inch material to 100 percent passing 6 mesh through a gyratory crusher, screening repeatedly.
- 4) Split out 3 Kg of the minus 6 mesh material and set up leach test.
- 5) Split out a 500 gram portion from the remaining minus 6 mesh material, pulverize, and send out for fire assay.

FIGURE 2. AURORA BUCKET LEACH TESTS
HEAD ASSAYS AND RECOVERY STATISTICS

| | | | | | | PERCENT | | | CHEMICAL | CONSUMPTION | FINENESS |
|-------------|---------------|------------------|------------------|------------------------|--------------------|-----------------|---------------|--------------|---------------|----------------------------------|--------------------|
| TEST NO. | SAMPLE NO. | SIZE (Inches) | DAYS LEACHING | oz Au/Ton RECOVERED | oz Au/Ton TAILS | Au RECOVERED | HEAD ASSAY | CALC HEAD | 1bs/S NaCN | Short Ton Ca(OH) ₂ | RECOVERED METAL |
| 1702 | 1701 A | 1-1/2 | 112 | .153 | .213 | 41.80 | | .366 | 5.16 | 1.01 | 667 |
| 1703 | 1701 B | 1-1/2 | 112 | .065 | .054 | 54.62 | .105 | .119 | 5.18 | 1.18 | 659 |
| 1704 | 1701 D | 1-1/2 | 112 | .034 | .020 | 62.96 | | .054 | 5.55 | 1.45 | 478 |
| 1736 | 1701 C | - 6M | 56 | .086 | .023 | 78.90 | .105 | .109 | 13.95 | 0.76 | 447 |
| 1737 | 1701 C | 5/8 | 56 | .085 | .031 | 73.28 | .105 | .116 | 10.70 | 0.71 | 514 |
| 1769 | 1759 A | 2 | 89 | .101 | .042 | 70.63 | .199 | .143 | 4.61 | 0.89 | 341 |
| 1770 | 1759 A | 5/8 | 89 | .137 | .049 | 73.66 | .199 | .186 | 4.53 | 0.74 | 311 |
| 1771 | 1759 B | 2 | 89 | .008 | .003 | 72.73 | .010 | .011 | 6.61 | 1.00 | 609 |
| 1772 | 1759 В | 5/8 | 89 | .007 | .002 | 77.78 | .010 | .009 | 5.35 | 1.58 | 521 |
| 1773 | 1759 C | 2 | 89 | .056 | .042 | 57.14 | .079 | .098 | 3.46 | 0.79 | 643 |
| 1774 | 1759 C | 5/8 | 89 | .062 | .023 | 72.94 | .079 | .085 | 3.07 | 1.49 | 624 |
| 1795 | 1759 D | 5/8 | 90 | .057 | .026 | 68.67 | .081 | .083 | 5.78 | 0.68 | 602 |
| 1959 | 1902 | 2 | 62 | .199 | .196 | 50.38 | .412 | .395 | 3.92 | 0.80 | 431 |
| 1960 | 1902 | 5/8 | 62 | .270 | .158 | 63.08 | .412 | .428 | 3.41 | 0.76 | 615 |
| | | | | | | | | | | | |

Samples 1759 A - C were treated as follows:

- 1) Crush entire sample to 2-inches through jaw crusher.
- 2) Split out a 5 gallon bucket of 2-inch material and set up bucket leach test.
- 3) Crush remaining 2-inch material to 5/8-inch through jaw crusher. Split out a 5 gallon bucket of 5/8-inch material and set up bucket leach test.
- 4) Take remaining 5/8-inch material and split out a 5 Kg sample. Crush to 100 percent passing 6 mesh through a gyratory crusher, screening repeatedly.
- 5) Split out two 500 gram portions from the minus 6 mesh material, pulverize, and submit for fire assay.
- 6) Run 24 hour cyanide bottle roll tests on pulverized and minus 6 mesh material.
- 7) Run 1 hour and 24 hour cyanide centrifuge tests on pulverized samples.

Samples 1759 D and E were treated as follows:

- 1) Crush entire sample to 5/8-inch through jaw crusher.
- 2) Split out 5 Kg of 5/8-inch material and crush to 100 percent passing 6 mesh through a gyratory crusher, screening repeatedly.
- 3) Split out two 500 gram portions from the minus 6 mesh material, pulverize, run centrifuge tests and send out for fire assay.
- 4) Take remaining 5/8-inch material from 1759 D only, split out a 5 gallon bucket, and set up a bucket leach test.
- 5) Run 24 hour cyanide bottle roll tests on pulverized and minus 6 mesh material from 1759 D.

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Sample 1902 was treated and prepared the same as samples 1759 A - C, however, only one hour centrifuge tests were run on the pulverized head sample. No bottle roll tests were conducted.

CYANIDE BOTTLE ROLL TESTS

Cyanide bottle roll tests were conducted on pulverized and minus 6 mesh head samples from samples $1759 \ A - D$, according to the following procedure:

- 1) Weigh out 100 grams of ore and place into 250 ml polybottle.
- 2) Add 150 mls of water and adjust pH to 10 using lime. Add 0.75 grams NaCN (equivalent to 5 gpl NaCN).
- 3) Place on rolls for 24 hours.
- 4) Filter; dry tailings and save.
- 5) Check solution for pH, Au, Ag and Cu.

Figure 3 shows the percent recovery of contained gold in a 24 hour cyanide bottle roll test (based on calculated head assays), and the product fineness (ratio of gold to gold plus silver; times 1000). On the average, 93 percent of contained gold was recovered in the tests on pulverized samples. Tests on unpulverized samples averaged 63 percent of contained gold recovered.

CENTRIFUGE TESTS ON HEAD SAMPLES

The pulverized pulps from head samples were subjected to cyanide centrifuge tests according to the following procedures:

- Weigh out 10 grams of pulverized ore and place in centrifuge tube.
- Add 25 mls of 5 gpl NaCN solution. Adjust pH, if necessary, to pH 10, using lime.

FIGURE 3. 24-HR CYANIDE BOTTLE ROLL TESTS
ON PULVERIZED AND MINUS 6 MESH SAMPLES

| SAMPLE NO. | TEST NO. | SIZE | oz Au/Ton RECOVERED | PERCENT RECOVERY | FINENESS RECOVERED METAL |
|---------------|-------------|------|------------------------|---------------------|-----------------------------|
| | | | | | Yalahay |
| 1759 A | 1793 A | P | .184 | 92.5 | 130 |
| 1759 A | 1793 E | -6M | .110 | 55.3 | 120 |
| 1759 B | 1793 В | P | .009 | 81.8 | 317 |
| 1759 В | 1793 F | -6M | .008 | 72.7 | 432 |
| 1759 C | 1793 C | P | .082 | 103.8 | 487 |
| 1759 C | 1793 G | -6M | .053 | 67.1 | 550 |
| 1759 D | 1793 D | P | .077 | 95.1 | 454 |
| 1759 D | 1793 Н | -6M | .044 | 54.3 | 469 |
| | | | | | |

FIGURE 4. AGITATED CYANIDE CENTRIFUGE TESTS
ON PULVERIZED PORTIONS OF SAMPLE

| SAMPLE NO. | TEST NO. | LEACH TIME HOURS | oz Au/Ton RECOVERED | PERCENT RECOVERY | FINENESS RECOVERED METAL |
|---------------|-------------|---------------------|------------------------|---------------------|-----------------------------|
| 1759 A | 1776 A | 1 | .131 | 65.8 | 109 |
| 1759 A | 1776 F | 24 | .201 | 101.0 | 136 |
| 1759 В | 1776 B | 1 | .013 | 118.2 | 310 |
| 1759 В | 1776 G | 24 | .009 | 81.8 | 235 |
| 1759 C | 1776 C | 1 | .086 | 108.9 | 484 |
| 1759 C | 1776 Н | 24 | .089 | 112.7 | 473 |
| 1759 D | 1776 D | 1 | .074 | 91.4 | 440 |
| 1759 D | 1776 I | 24 | .080 | 98.8 | 445 |
| 1759 E | 1776 E | 1 | .050 | 56.2 | 358 |
| 1759 E | 1776 J | 24 | .049 | 55.6 | 354 |
| 1902 | 1915 B | 1 | .274 | 66.5 | 303 |
| | | | | | |

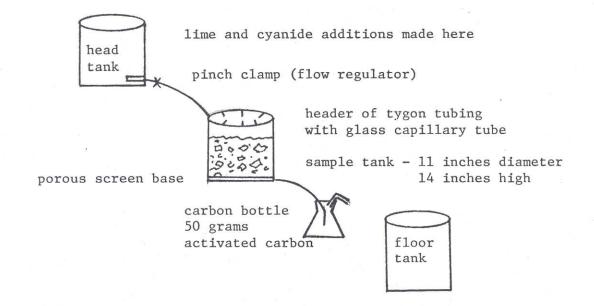
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- 3) Place on wrist-action shaker for specified time.
- 4) Centrifuge and filter through glass wool.
- 5) Check solution for pH, Au, Ag and discard residue.

Figure 4 shows the percent recovery of contained gold in cyanide centrifuge tests on pulverized material, and the product fineness (ratio of gold to gold plus silver; times 1000).

BUCKET LEACH TEST APPARATUS

The apparatus for the 2, 1-1/2 and 5/8-inch rock tests is shown in the drawing below. The apparatus for tests 1736 and 1737 differed slightly, in that the sample tank was 4-inches in diameter and 18-inches high.



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LEACH TEST PROCEDURE

In the apparatus shown on the preceding page, the center tank, or leach tank, was filled with the rock to be leached.

Alkaline cyanide solution was continuously distributed onto the rock from the head tank, through a set of glass capillary drip tubes. Flowrate of solution dripping onto the rock was controlled using a pinch clamp.

Solutions entering the floor tank were assayed every two cycles for cyanide and lime and reagents were added, as necessary, to maintain solutions at "target levels".

Solutions exiting the leach tank flowed continuously through a bottle of activated carbon and then into a floor tank. The active solution in the system was recycled to the head tank every 48 to 72 hours.

The tanks were kept covered at all times to minimize evaporation and cyanide loss. No makeup water was required.

The charge of activated carbon was removed three times during the tests, with the exception of tests 1702, 1703 and 1704, which had five carbon changes, and assayed to determine the amount of gold and silver leached from the ore.

TEST HISTORIES

Start-up of Tests. The initial leach solution for all tests contained 1.0 grams NaCN per liter and 0.5 grams Ca(OH) per liter. Initial solutions exiting all tests were alkaline and contained measureable amounts of cyanide.

 $\frac{\hbox{Solution Color and Clarity.}}{\hbox{brown throughout the tests.}} \quad \hbox{Solutions remained clear and slightly}$

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Cyanide Strength and Alkalinity. Cyanide strength was allowed to decline slowly for the first 14 days of the tests to a minimum of 0.2 to 0.5 grams NaCN per liter. It was then maintained in the range 0.4 to 0.6 grams per liter for the remainder of the tests.

Alkalinity was generally maintained, with lime, in the range pH 10.0 to 10.6 for the duration of all tests. Lime and cyanide consumption data are tabulated in Figure 2.

Tests 1702, 1703 and 1704 were started, using as-received ore, with a maximum dimension of 4-inches. Initial gold recovery was poor with 10 percent of contained gold (17.5 percent of recoverable gold) being recovered onto carbon by day 14. The tests were stopped on day 15, the ore allowed to air dry, then crushed to 1-1/2-inches and the tests restarted. Analysis of the leach solutions by atomic absorption showed that after 7 days leaching at 1-1/2-inches, an additional 25 percent of contained gold (46 percent of recoverable gold) had been leached, on average.

GOLD AND SILVER RECOVERIES

Figure 2 tabulates gold recoveries and fineness of recovered metal (ratio of gold to gold plus silver; times 1000). Figures 5 and 6 show gold recovery based on carbon assays, versus days leaching.

Average recovery for all tests was 64 percent of contained gold. Recoveries from tests on ore crushed to 2-inches and 1-1/2-inches (7 tests), averaged 59 percent of contained gold. Recoveries from tests on ore crushed to 5/8-inch (6 tests), averaged 72 percent of contained gold. The average fineness of the beads obtained from fire assay of the activated carbon was 540.

TAILINGS ASSAY AND METALLURGICAL BALANCES

At the end of the tests, the test tailings were dried and then screened into various size fractions. The size fractions were crushed, if necessary, to 100 percent passing 6 mesh and then two portions were split out and pulverized. Fire assays were run on each of the pulverized portions. Tailings assays and weights are reported in Figure 7.

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Metallurgical balances, or comparisons between original head assays and calculated head assays (gold recovered onto carbon + gold remaining in the tailings) were good for all tests, averaging 90.3 percent. Eight of the 12 tests showed higher calculated heads than assay heads, four showed the reverse; three showed calculated heads within .005 ounces per ton of the assay heads. The calculated head, based on actual recoveries and assays of the fine-crushed tailings, is considered more accurate than the assay head.

ASSAYING PROCEDURES

 $\underline{\text{Heads and Tailings Assays}}$. Heads and tailings assays were all run in duplicate, half as one assay ton fire assays and half as half assay ton fire assays.

<u>Carbon Assays.</u> The loaded activated carbon was dried and weighed. Two samples were split out and assayed and the remainder saved for reference. The carbon for assays was roasted to convert it to ash, then conventionally fire assayed.

Solution Assays. Approximate solution assays were made periodically on an atomic absorption spectrophotometer, using as a standard, a gold cyanide solution which had been calibrated by fire assay. The solution assays were used merely to check on the progress of the leach, since actual recovery was based on fire assay of the activated carbon.

Final solution was checked by AA methods and found to contain negligible amounts of gold.

Submitted by,

Russell B. Dix

RBD/df

APPENDIX B

STATISTICAL ANALYSIS OF SAMPLE DATA AND SENSITIVITY STUDY

In an attempt to arrive at a more realistic mean value of ore grade based on assay results within the ore zone defined by 1981 rotary drilling, Prof. Pierre Mousset-Jones of the Mackay School of Mines, Reno, Nevada, was asked to recalculate the mean values under several cases, using methods developed by H.S. Sichel (1966) for lognormally distributed gold assays. Using the School of Mines computer and software, M. Klemme, a mining engineering graduate student, developed the data presented in Figure 1.

Case III B, the lognormal mean determined after excluding the two highest assays and the two zero assays, probably represents the best mean value of the cases considered, as far as this method is able to yield an accurate estimate of the mean.

Professor Mousset-Jones suggested that due to probable assay value dependence (i.e., samples are not statistically speaking, truly independent), the indicated confidence limits are slightly optimistic. In otherwords, at the 90 percent confidence level, the upper limit may actually be a somewhat higher value and the lower limit may actually be somewhat lower than the values indicated.

It should be pointed out that although Sichel's method is less sophisticated than modern geostatistical methods, it probably yields a somewhat more realistic mean value of the ore zone than simple arithmetical averaging.

A complete description of the techniques for this study are in "An Introduction to Geostatical Methods of Mineral Evaluation" by J.M. Rendu, Chapter 2, a copy of which is at the end of this report.

STATISTICAL ANALYSIS

A total of 142 samples (103 rotary drillhole and 39 rock chip samples) taken from the designated ore zone (Block A) in 1981, were used in the study. The different cases examined were as follows:

| CASE NUMBER | CONDITIONS |
|----------------|---|
| I A | Normal Distribution. All samples included (N = 142). |
| II A | Normal Distribution. All samples excluding zero values (N = 140). |
| II B | Lognormal Distribution. All samples excluding zero values ($N = 140$). |
| III A | Normal Distribution. All samples excluding two highest assays and zero values ($N = 138$). |
| III B | Lognormal Distribution. All samples excluding two highest assays and zero values ($N = 138$). |

The results of these analyses are presented in Figure 1. Figures 2 and 3 list the samples used for the study along with their fire assays. Figure 4 presents the cumulative frequency of the samples. A class size of 0.030 ounces per ton gold was used for the entire study.

Case III B, the worst possible case, indicates that there is a 90 percent chance that the average ore grade lies between the upper limit of 0.1480 ounces per ton gold and the lower limit of 0.0980 ounces per ton, with a 5 percent chance it may lie below 0.0980 ounces per ton.

SENSITIVITY ANALYSIS

Figure 5 presents a sensitivity analysis of the feasibility study to variations of grade, percent recovery and gold prices. Using the minimum ore grade obtained from the statistical analysis (Case III B) and a gold price of \$360.00 per ounce, the venture will still have higher returns than costs.

An alternative sensitivity scenario analysis is also presented on the effects on costs if the same total ounces of gold are recovered as originally expected, at a price of \$360.00 per ounce, if due to dilution, an additional 20 percent by weight of "waste" rock (11,000 tons) is diverted to the heaps to be crushed and leached. The venture shows a positive return here also.

Submitted by,

Russell B. Dix

RBD/df

FIGURE 1
SUMMARY OF ORE GRADE CALCULATIONS

| CASE NUMBER | MEAN (u) | 90% CONFIDENCE UPPER | LIMITS ¹ LOWER |
|----------------|----------|-------------------------|---------------------------|
| IA | 0.1229 | 0.1428 | 0.1030 |
| II A | 0.1247 | 0.1448 | 0.1046 |
| II B | 0.1260 | 0.1580 | 0.1050 |
| III A | 0.1149 | 0.1315 | 0.0983 |
| III B | 0.1180 | 0.1480 | 0.0980 |

¹ - Values in ounces gold per short ton of ore.

FIGURE 2

AURORA DRILLHOLE SAMPLES USED

FOR ORE GRADE STATISTICS

| DRILLHOLE NUMBER | INTERVAL (Feet) | FIRE ASSAY ounces/ton Au | DRILLHOLE NUMBER | INTERVAL (Feet) | FIRE ASSAY ounces/ton Au |
|---------------------|--------------------|-----------------------------|---------------------|----------------------------|-----------------------------|
| PH-15 | | 0.119 | PH-49 | 15- 20 | 0.042 |
| | | 0.017 | | 20- 25 | 0.014 |
| | | | | 25- 30 | 0.014 |
| PH-39 | 45- 50 50- 55 | 0.046 0.036 | PH-51 | 40- 45 | 0.096 |
| | 55- 60 | 0.080 | | 45- 50 | 0.380 |
| | 60- 65 | 0.086 | | 50- 55 | 0.272 |
| ±∈ | 65- 70 | 0.086 | | 55- 60 | 0.164 |
| | 70 - 75 | 0.057 | | 60- 65 | 0.064 |
| | 75 - 80 | 0.074 | | 65- 70 | 0.122 |
| | 80- 85 | 0.053 | | 70- 75 | 0.234 |
| | 85- 90 | 0.560 | ¥ | 75- 80 | 0.140 |
| | | | | 80- 85 | 0.312 |
| PH-40 | 20- 25 | 0.049 | | 85- 90 | 0.100 |
| | 25- 30 | 0.028 | | 90- 95 | 0.086 |
| | 30- 35 | 0.049 | | 95-100 | 0.240 |
| | 35- 40 | 0.024 | | 100-105 | 0.142 |
| | 40- 45 | 0.260 | | 105-110 | 0.191 |
| | 45- 50 | 0.051 | | 110-115 | 0.408 |
| | 50- 55 | 0.024 | | 115-120 | 0.902 |
| PH-43 | 45- 50 | 0.041 | PH-52 | 5- 10 | 0.077 |
| 70377 NATA | 50- 55 | 0.061 | Y. II | 10- 15 | 0.098 |
| | 55- 60 | 0.070 | n e | 15- 20 | 0.081 |
| | 60- 65 | 0.054 | | 20- 25 | 0.017 |
| | | | | 25- 30 | 0.163 |
| PH-44 | 5- 10 | 0.162 | | 30- 35 | 0.317 |
| | 10- 15 | 0.126 | | 35- 40 | 0.133 |
| | 15- 20 | 0.040 | | 40- 45 | 0.064 |
| | 20- 25 | 0.040 | | 45- 50 | 0.126 |
| | 25- 30 | 0.030 | | 50- 55 | 0.038 |
| | 30- 35 | 0.030 | PH-53 | 50- 55 | 0.060 |
| PH-45 | 40 45 | 0.001 | 111 55 | 55- 60 | 0.048 |
| PH-43 | 40- 45 45- 50 | 0.081 | | 60- 65 | 0.174 |
| | 50- 55 | 0.112 0.012 | | 65- 70 | 0.010 |
| | 55- 60 | 0.012 | | 70- 75 | 0.012 |
| | 60- 65 | 0.033 | | 75- 80 | 0.451 |
| | 65- 70 | 0.288 | | 80- 85 | 0.145 |
| | 70- 75 | 0.053 | | | |
| | 75- 80 | 0.073 | PH-60 | 45- 50 | 0.520 |
| | 80- 85 | 0.031 | | 50- 55 | 0.044 |
| | 85- 90 | 0.085 | | 55- 60 | 0.156 |
| | 90- 95 | 0.100 | | 60- 65 | 0.400 |
| | 95-100 | 0.206 | | 92179.W 1254.9 x | |
| | | | DDH-2 | 65- 70 | 0.124 |
| PH-47 | 10- 15 | 0.063 | | 70- 78 | 0.258 |
| | 15- 20 | 0.057 | | 78- 82 | 0.038 |
| | 20- 25 | 0.015 | | 82- 87 117 - 122 | 0.122 |
| | 25- 30 | 0.026 | | 122-127 | 0.040 |
| | 30- 35 35- 40 | 0.038 | | 127-127 | 0.028 |
| | 40- 45 | 0.048 0.109 | | 131-136 | 0.368 |
| | 45- 50 | 0.015 | | 136-142 | 0.366 |
| PH-48 | 5- 10 | 0.032 | | | |
| rn-40 | 10- 15 | 0.025 | | | |
| | 15- 20 | 0.135 | 2.8 | | |
| | 20- 25 | 0.031 | | | |
| | 25- 30 | 0.032 | | | |
| | 30- 35 | 0.015 | | | |

FIGURE 3
OUTCROP CHIP SAMPLES USED FOR STATISTICAL ANALYSIS

| ORE BLOCK A SECTION | LINE | FIRE ASSAY Au oz/ton |
|---------------------|-----------|-------------------------|
| 1 | Northern | .062 |
| | | .095 |
| | | .132 |
| | | .029 |
| 1 | Center | .038 |
| ** | 0011002 | .545 |
| | | .286 |
| | | .167 |
| | | .063 |
| 1 | Southern | .029 |
| 1 | Southern | |
| | | .037 |
| | | .049 |
| * . | | .148 |
| | | .075 |
| | 8 | .057 |
| | | .022 |
| | | .013 |
| | | .008 |
| | | .003 |
| 2 | Adit | .000 |
| | 555 W 35 | .000 |
| | | .155 |
| | | .097 |
| 3 | Northern | .108 |
| 3 | | .401 |
| | | .088 |
| | | |
| 3 | Center | .180 |
| 3 | Southern | .119 |
| | | .032 |
| | | .180 |
| | | .053 |
| | | .020 |
| 4 | Northern | .061 |
| *** | | .159 |
| 4 | Southern | .057 |
| 4 | Doublieth | .072 |
| | | .695 |
| | | .142 |
| | | |
| | | .047 |

FIGURE 4

CUMULATIVE FREQUENCY

DISTRIBUTION OF SAMPLE VALUES

(Class Interval = 0.03)

| CELL NUMBER | CLASS UPPER VALUE | NUMBER OF SAMPLES | CUMULATIVE SAMPLES | CUMULATIVE FREQUENCY (%) |
|----------------|----------------------|----------------------|-----------------------|-----------------------------|
| 1 | 0.030 | 10 | 10 | 7.0 |
| 2 | 0.060 | 36 | 46 | 32.4 |
| 3 | 0.090 | 29 | 75 | 52.8 |
| 4 | 0.120 | 15 | 90 | 63.4 |
| 5 | 0.150 | 12 | 102 | 71.8 |
| 6 | 0.180 | 12 | 114 | 80.3 |
| 7 | 0.210 | 5 | 119 | 83.8 |
| 8 | 0.240 | 1 | 120 | 84.5 |
| 9 | 0.270 | 2 | 122 | 85.9 |
| 10 | 0.300 | 3 | 125 | 88.0 |
| 11 | 0.330 | 3 | 128 | 90.1 |
| 12 | 0.360 | 1 | 129 | 90.8 |
| 13 | 0.390 | 2 | 131 | 92.2 |
| 14 | 0.420 | 3 | 134 | 94.4 |
| 15 | 0.450 | 1 | 135 | 95.1 |
| 16 | 0.480 | 2 | 137 | 96.5 |
| 17 | 0.510 | 0 | 137 | 96.5 |
| 18 | 0.540 | 1 | 138 | 97.2 |
| 19 | 0.570 | 1 | 139 | 97.9 |
| 20 | 0.600 | 1 | 140 | 98.6 |
| 21-23 | 0.690 | 0 | 140 | 98.6 |
| 24 | 0.720 | 1 | 141 | 99.3 |
| 25-30 | 0.900 | 0 | 141 | 99.3 |
| 31 | 0.930 | 1 | 142 | 100.0 |

FIGURE 5

- SENSITIVITY OF AURORA FEASIBILITY STUDY TO VARIATIONS OF GRADE, PERCENT RECOVERY AND GOLD PRICES

At .098 oz/ton Au (Case III b Lower 90% Confidence Limit) for Ore Block A and .045 oz/ton Au for Ore Block B at 65% Recovery

 $.098 \text{ oz/ton } \times 41,000 \text{ tons} = 4018 \text{ oz}$

.045 oz/ton x 15,000 tons = 675 oz

TOTAL

4693 oz

 $4693 \text{ oz } \times .65 = 3050 \text{ oz recoverable gold}$

3050 oz @ \$400.00 = \$1,220,180.00

@ \$360.00 = \$1,098,000.00

- 1,091,000.00

- 1,091,000.00

PROFIT

\$ 129,180.00

\$ 7,000.00

At .116 oz/ton Au @ \$360.00 Gold and 65% Recovery

.116 oz/ton x 41,000 tons = 4756 oz

5506 oz x .65 = 3579 oz recovered

.05 $oz/ton \times 15,000 tons = 750 oz$

TOTAL

5506 oz

 $3579 \text{ oz } \times \$360.00 = \$1,288,440.00$

\$ 1,288,440.00

- 1,091,000.00

PROFIT

\$ 197,440.00

At .116 oz/ton Au @ \$360.00 Gold, 65% Recovery and an additional 20% of Waste diverted to the heaps

.116 oz/ton x 41,000 tons = 4756 oz

5506 oz x .65 = 3579 oz recovered

.050 oz/ton x 15,000 tons = 750 oz

.000 oz/ton x 11,000 tons = 0 oz

TOTAL

5506 oz

 $3579 \text{ oz } \times \$400.00 = \$1,431,600.00$

@ \$360.00 = \$1,288,440.00

- 1,162,500.00

-1,162,500.00

PROFIT

\$ 269,100.00

\$ 125,940.00

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At 70% Recovery vs/65% Recovery, Same Grade

41,000 tons @ .116 oz/ton = 4756 oz 5506 oz x .70 = 3854 oz recovered

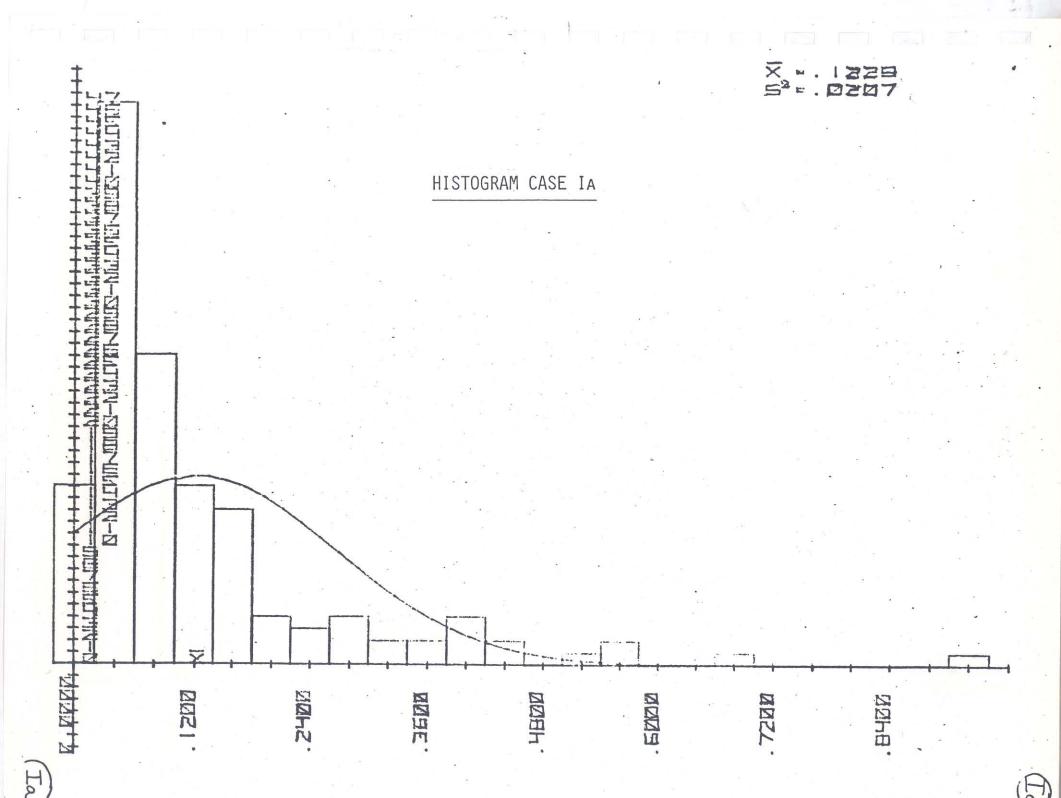
15,000 tons @ .05 oz/ton = 750 oz

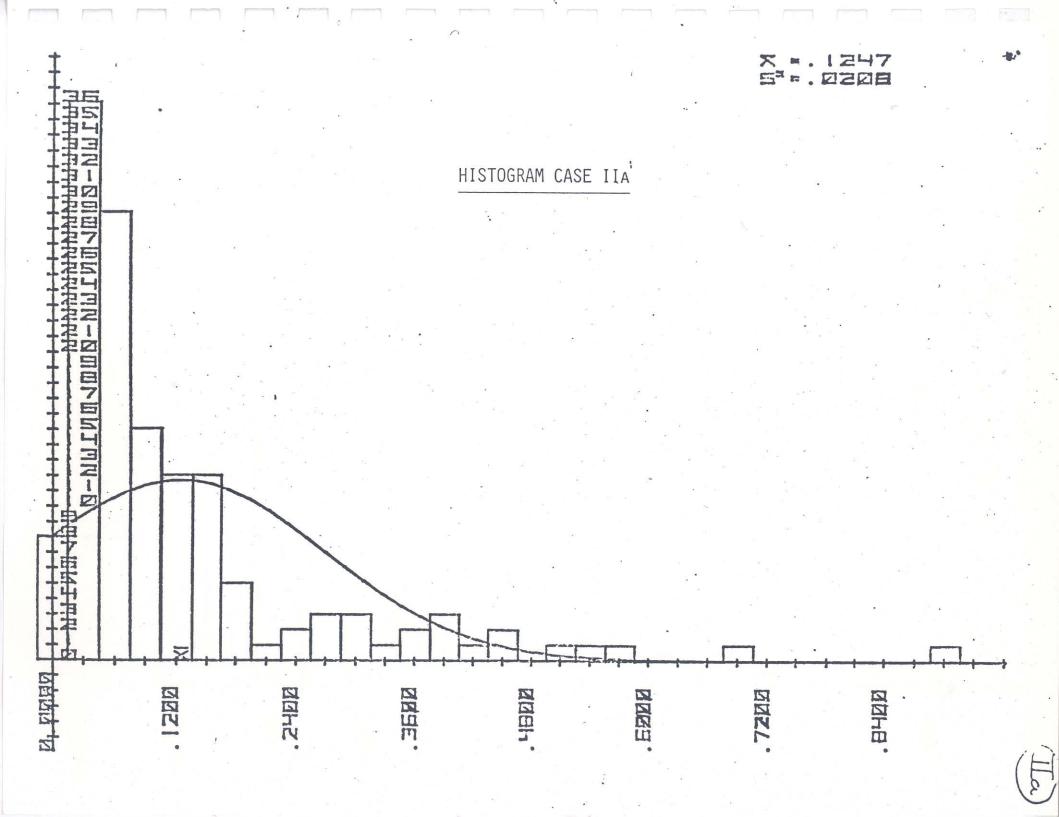
TOTAL 5506 oz

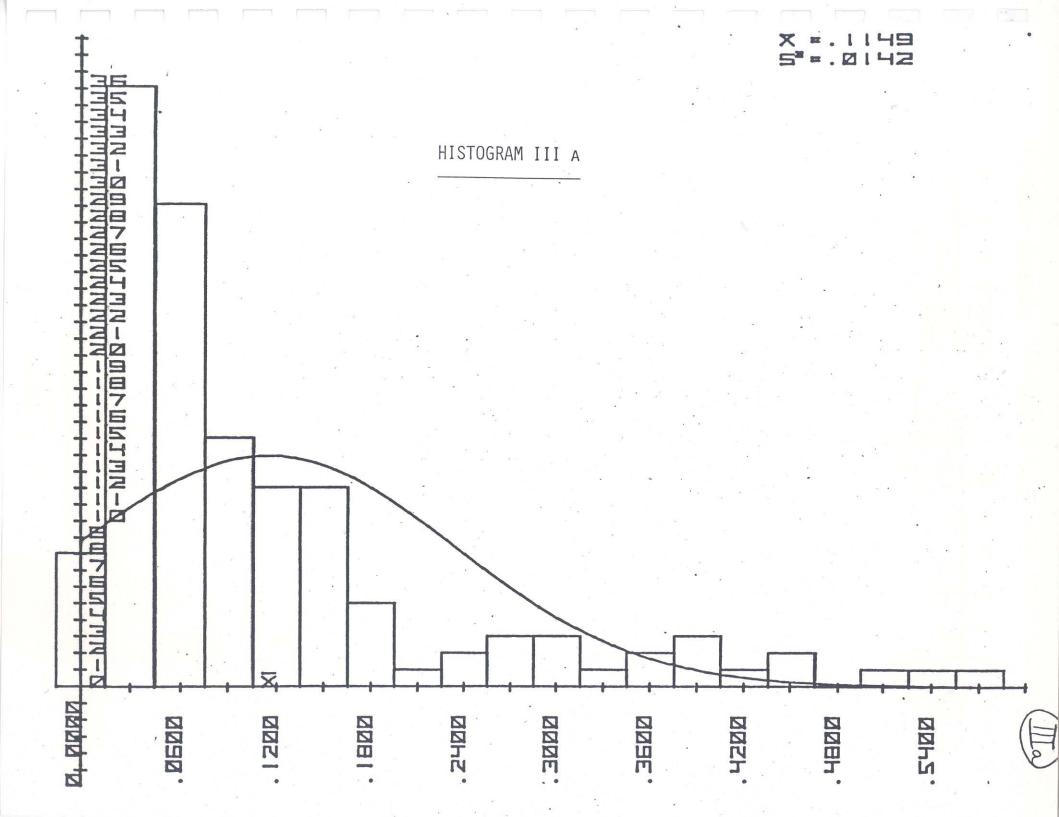
 $3854 \text{ oz } \times \$400.00 = \$1,541,680.00$ 3854 oz @ \$360.00 = \$1,387,440.00

- 1,091,000.00 - 1,091,000.00

PROFIT \$ 450,680.00 \$ 296,440.00







CUMULATIVE FREQUENCY KOE PROBABILITY 46 8043 99.9 99.8 0 0 NORMAL -09-008 -07_ -O 6 .04 .03 .02

An Introduction to Geostatistical Methods of Mineral Evaluation

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Johannesburg 1978

CHAPTER 2

Classical Statistics, Random Distributions, Normal and Lognormal Theory

2.1 General

To determine the characteristics of a mineral deposit, the usual practice is to take samples, analyse the properties of those samples and infer the characteristics of the deposit from these properties. This analysis can be done using statistical methods. In the present work we are concerned with two types of statistical approaches, classical statistics and spatial statistics.

If one uses classical statistics to represent the properties of the sample values, the assumption is made that the values are realizations of a random variable. The relative positions of the samples are ignored, and it is assumed that all sample values in the mineral deposit have an equal probability of being selected. The likely presence of trends, zones of enrichment, or pay shoots in the mineralization, is ignored. The fact that two samples taken close to each other are more likely to have similar values than if taken far apart is also not taken into consideration.

In contrast, spatial statistics will be used if one chooses to consider that the sample values are realizations of random functions. On this hypothesis, the value of a sample is a function of its position in the mineralization of the deposit, and the relative position of the samples is taken into consideration. The similarity between sample values is quantified as a function of the distance between samples and this relationship represents the foundation of spatial statistics.

There are few situations where classical statistics can be used. The assumption that all sample values in the mineral deposit have an equal likelihood of being represented will be satisfied only if the sample values are randomly distributed, or if the sample positions are random. Sample values are in fact never randomly distributed within a mineral deposit. Furthermore, geologists usually avoid taking samples at random (random sampling), as it is correctly accepted that samples located on a regular grid, or approximately on a regular grid, give more information than randomly located samples. In practice, classical statistics should be used only in the early stages of exploration, when the number of samples available is relatively small and the distances between samples are large. In these circumstances, and whenever the information available is not sufficient to permit the use of spatial statistics, application of the methods described in this chapter is justified.

2.2 Definitions and remarks

In considering a mineral deposit, the symbol Ω will represent the domain (volume or surface) which constitutes this deposit. Consider a point z within Ω , and a sample w centred at point z. A value x(z) is associated with this sample. For example, x(z), which is a function of the position z where the sample is taken, can be the grade of ore, the thickness, the content per unit area, or any other quantity

characteristic of the sample. For this reason x(z) is called a regionalized variable. The value x(z) is a function of the size and orientation of the sample w, which is defined as the support of the regionalized variable (Fig. 2.1).



Fig. 2.1. Position z and value x(z) of a sample w in an ore body Ω .

If we take all possible samples w at all possible points z within the ore body Ω , we can calculate the average value μ of all the x(z) values in the ore body, and this value is independent of the support w. The following notation will be used:

$$\mu = E_{\Omega}[x(z)]$$
 expected value of $x(z)$ in Ω . (2.1)

In the early stages of exploration, the main problem of analysis is the estimation of μ . For this purpose, n samples of the same support w are taken at points z_i , i=1,2...n. The value of the ith sample is $x(z_i)$. The sample values are used to calculate an estimate $\hat{\mu}$ of the mean μ , and confidence limits for the mean (the symbol * will be used throughout this work to indicate an estimator, or the value of an estimator). The estimator for this purpose will vary according to the probability distribution of x(z). Since in this analysis it is assumed

ith sample can be ignored, and we can use the notation
$$x = x(z)$$
 and:
 $x_i = x(z_i)$. (2.2)

that all sample values are independent, the location zi of the

$$\mu = E[x]. \tag{2.3}$$

Only two types of sample value distribution will be considered here, the normal distribution and the lognormal distribution. Many other types of distribution can be found in the literature (Sichel, 1973; David, 1977; Becker, 1964–1966), but their study would go beyond the limits of our present requirements. In most practical situations the assumption of either normality or lognormality can be made, and the use of more complex distributions is not justified.

2.3 The normal distribution

2.3.1 General

Suppose we have n sample values x_i , i = 1, 2, ..., n. The first step in the analysis of these values consists in grouping them in classes, and in counting the number of samples which fall within each class. The possible result of such an analysis

Table 2.1

Calculation of the percent cumulative frequency distribution of sample values

| Column No. Class No. | 1 Sample value | 2 Seconds | 3 | 4 | 5 |
|-------------------------|-------------------|------------------------------|----------------------------|---|---------------|
| C1833 110. | lower limit (%) | Sample value upper limit (%) | Number of samples in class | Cumulative freq. (Number of samples) | Cum, freq |
| 1 | 5 | 10 | 1 | The state of the state of | 1,4 |
| 2 | 10 | 15 | 3 | 4 | 5,6 |
| 3 | 15 | 20 | 2 | 6 | 8,3 |
| 4 | 20 | 25 | 5 | 11 | 15,3 |
| 5 | 25 | 30 | 6 | 17 | |
| 6 | 30 | 35 | 9 | 26 | 23,6 |
| 7 | 35 | 40 | 11 | 37 | 36,1 |
| 8 | 40 | 45 | 10 | 47 | 51,4 |
| 9 | 45 | 50 | 7 | 54 | 65,3 |
| 10 | 50 | 55 | 8 | 62 | 75,0 |
| 11 | 55 | 60 | 5 | 67 | 86,1 |
| 12 | 60 | 65 | 3 | 70 | 93,1 |
| 13 | 65 | 70 | 0 | 70 | 97,2 |
| 14 | 70 | 75 | 1 | | 97,2 |
| 15 | 75 | 80 | i | 71 72 | 98,6 100,0 |

is given in Table 2.1 (columns 1, 2 and 3). Using this result it is possible to draw the histogram illustrated in Fig. 2.2. The histogram is a valuable tool in determining whether the sample distribution is reasonably symmetrical, and to detect visually possible outliers, or sample values which may be abnormally high or low. However, the shape of the histogram is affected by the limits of the classes used to group the samples.

To determine whether sample values are normally distributed a convenient graphical method can be used. One calculates the cumulative frequency distribution of the values (Table 2.1, columns 4 and 5) and plots this distribution on normal probability paper (Fig. 2.3). From the definition of the normal probability scale, the cumulative distribution of a normally distributed variable will be plotted as a straight line on normal probability paper. If the points obtained by plotting the cumulative distribution of the sample values in this way can be considered as distributed along a straight line, the assumption of normal distribution can be accepted, and the theory described in this section can be used.

In practice, the assumption of normal distribution of the sample values is not often satisfied except when the sampled mineralization has a relatively high grade (for example, some iron ore deposits) or when the value considered has a low variability (for example, the average thickness of some bedded deposits, or the specific gravity of most mineralizations). When taking samples from a new ore body, the number of samples might be insufficient to obtain a representative histogram or cumulative distribution. Judgment and past experience are then used to decide whether the assumption of normality can be accepted.

2.3.2 Mean, variance and confidence intervals estimation

The sample mean and sample variance are estimated as follows:

Sample mean
$$\bar{x} = \frac{1}{n} \sum_{i=1}^{n} x_i$$
. (2.4)

Sample variance
$$s^2 = \frac{1}{n-1} \sum_{i=1}^{n} (x_i - \bar{x})^2$$
, (2.5)

or
$$s^{2} = \frac{1}{n-1} \left(\sum_{i=1}^{n} x_{i}^{2} - n(\bar{x})^{2} \right). \tag{2.6}$$

where $s = \sqrt{s^2}$ is the estimator of the standard deviation of the sample population. The mean value of the ore body is estimated by:

$$\hat{\mu} = \bar{x}$$
with variance (2.7)

$$V(\mathring{\mu}) = \frac{s^2}{n}.$$
 (2.8)

Let μ_p be the confidence limit of the true mean μ such that the probability that μ is smaller than μ_p is p. Then μ_{1-p} is the confidence limit such that the probability that μ is larger than μ_{1-p} is p. The probability that μ falls between μ_p and μ_{1-p} is 1-2p and these limits are called the central 1-2p confidence limits of the mean. If n is greater than 25, the following approximation equations can be used to calculate central 68% and 95% confidence limits for the mean value of the ore body (p=16% and p=2,5% respectively):

central 68% confidence limits:
$$\bar{x} - \frac{s}{\sqrt{n}}$$
, $\bar{x} + \frac{s}{\sqrt{n}}$. (2.9)

central 95% confidence limits:
$$\bar{x} - 2\frac{s}{\sqrt{n}}$$
, $\bar{x} + 2\frac{s}{\sqrt{n}}$. (2.10)

If n is smaller than 25, μ_p and μ_{1-p} must be calculated as follows:

lower limit
$$\mu_p = \hat{\mu} - t_{1-p} \frac{s}{\sqrt{n}}$$
 (2.11)

upper limit
$$\mu_{1-p} = \hat{\mu} + t_{1-p} \frac{s}{\sqrt{n}}$$
 (2.12)

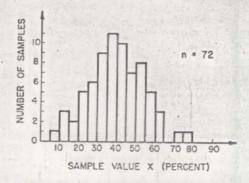


Fig. 2.2. Histogram of a symmetrical sample value distribution.

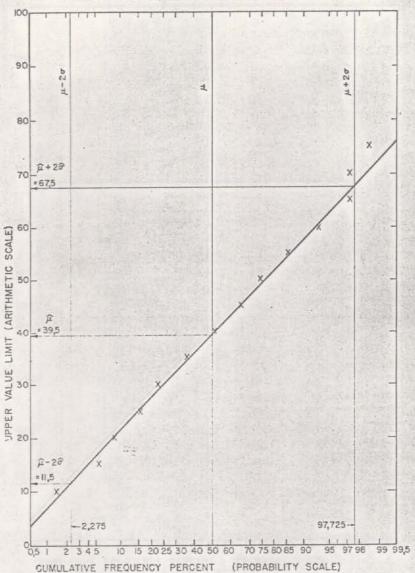


Fig. 2.3. Cumulative frequency distribution of a normal variate.

where t_{1-p} is the value of the Student's *t*-variate for f = n-1 degrees of freedom, such that the probability that t is smaller than t_{1-p} is 1-p. This value can be obtained in tables available in all textbooks on statistics (for example, Fraser, 1958, p. 389). For convenience, some values of t_p are reproduced in Table 2.2.

Example: For p=5% and n=10 we calculate 1-p=95% and f=n-1=9. Hence, using Table 2.2: $t_p=-t_{1-p}=-t_1833$.

2.3.3 Graphical estimation of mean and standard deviation

Provided the number of samples is large enough, the mean and standard deviation of the sample population can be estimated graphically (Fig. 2.3). Let x be a normal variate with mean μ and variance σ^2 . One needs only to remember the following probabilities:

Prob
$$(x < \mu) = 50\%$$

Prob $(x < \mu - 2\sigma) = 0,02275$
Prob $(x < \mu - 2\sigma) = 0,97725$.

For example, from Fig. 2.3 we estimate:

$$\hat{\mu} = 39.5\%$$
 $\hat{\sigma} = (67.5\% - 11.5\%) / 4 = 14.0\%.$

Exact confidence limits for μ cannot be obtained from these estimates. However, in practice, (2.11) and (2.12) above will be used, in which $\hat{\sigma}$ is substituted for s.

2.3.4 Example

Eight samples have been taken from an ore body, whose values are believed to be normally distributed. The sample values are given in Table 2.3.

- Draw the histogram, using class intervals of length 0,2.
 This is done on Fig. 2.4.
- (2) Estimate the average grade of the ore body and the corresponding central 90% confidence limits. This is

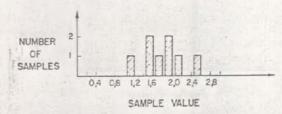


Fig. 2.4. Histogram of the sample values given in Table 2.3.

Table 2.2
Fractiles t_p of the t-distribution

| p | 000 | 2004 | | |
|---------|--------|-------|----------|-------|
| F = n-1 | 80% | 90% | 95% | 97,5% |
| 1 | 1,376 | 3,078 | 6,134 | 12,06 |
| 2 | 1,061 | 1,886 | 2,920 | 4,303 |
| . 3 | 0,978 | 1,638 | 2,353 | 3,182 |
| 4 | 0,941 | 1,533 | 2,132 | 2,776 |
| 5 | 0,920 | 1,476 | 2,015 | 2,571 |
| - 6 | 0,906 | 1,440 | 1,943 | 2,447 |
| 7 | 0,896 | 1,415 | 1,895 | 2,365 |
| 8 | 0,889 | 1,397 | 1,860 | 2,306 |
| 9 | 0,883 | 1,383 | 1,833 | 2,262 |
| 10 | 0,879 | 1,372 | 1,812 | 2,228 |
| 11 | 0,876 | 1,363 | 1,796 | 2,201 |
| 12 | 0,873 | 1,356 | 1,782 | 2,179 |
| 13 | 0,870 | 1,350 | 1,771 | 2,160 |
| 14 | 0,868 | 1,345 | 1,761 | 2,145 |
| 15 | 0,866 | 1,341 | 1,753 | 2,131 |
| 16 | 0,865 | 1,337 | 1,746 | 2,120 |
| 17 | 0,863 | 1,333 | 1,740 | 2,110 |
| 18 | 0,862 | 1,330 | 1,734 | 2,101 |
| 19 | 0,861 | 1,328 | 1,729 | 2,093 |
| 20 | 0,860 | 1,325 | 1,725 | 2,086 |
| 21 | 0,859 | 1,323 | 1,721 | 2,080 |
| 22 | 0,858 | 1,321 | 1,717 | 2,074 |
| 23 | 0,858 | 1,319 | 1,714 | 2,069 |
| 24 | 0,857 | 1,318 | 1,711 | 2,064 |
| 25 | 0,856 | 1,316 | 1,708 | 2,060 |
| 30 | 0,854 | 1,310 | 1,697 | 2,042 |
| 40 | 0,851 | 1,303 | 1,684 | 2,021 |
| 50 | 0,849 | 1,298 | 1,676 | 2,009 |
| 100 | 0,845 | 1,290 | 1,660 | 1,984 |
| | 0.0.00 | | 4 (1964) | |

(Reproduced from Fisher and Yates, Statistical Tables for Biological, Agricultural and Medical Research, Longman, 1974, with the permission of the publishers)

1,282

1,645

1,960

done using Table 2.4. The expected average value of the ore body is $\hat{\mu} = \bar{x} = 1,812$. Given 1-2p = 0,90 we calculate p = 0,05, and for n = 8 we derive (from Table 2.2) $t_{0,85} = 1,895$.

Hence the 95% confidence limits for the mean are:

Upper limit:
$$\mu_{1-p} = 1.812 + 1.895 \frac{0.402}{\sqrt{8}} = 2.08$$
.

Lower limit:
$$\mu_p = 1,812 - 1,895 \frac{0,402}{\sqrt{8}} = 1,54$$

1

1,2

2,0

1,6

Table 2.3
Sample values

2 3 4 5 6 7

2,5

1,9

1,5

1,7

There is a 90% chance that the average value of μ will be between 2,08 and 1,54, and a 5% chance that it be less than 1.54

Table 2.4
Calculation of mean and variance

| 1 | x_l | x_l^2 |
|-------------|-------|---------|
| i | 1,2 | 1,4 |
| 2 | 2,0 | 4,00 |
| 2 3 | 1,6 | 2,56 |
| | 1,7 | 2,89 |
| 4 5 6 | 2,5 | 6,25 |
| | 1,9 | 3,61 |
| 7 | 1,5 | 2,25 |
| 8 | 2,1 | 4,41 |
| Sum | 14,5 | 27,41 |

2.4 The lognormal distribution

2.4.1 General

In most situations, and certainly in the case of low-grade deposits, the distribution of the sample values is not symmetrical but positively skew (Fig. 2.5). These skew distributions can generally be well-represented by the two-parameter or three-parameter lognormal distribution. Let x be a variate with a skew distribution. If $\log_e(x)$ is a variate with normal distribution, then the distribution of x is said to be two-parameter lognormal. If $\log_e(x+\beta)$ is a normal variate, where β is a constant, then x is a three-parameter lognormal variate. The effect of the additive constant β on the histogram of a three-parameter lognormal variate is shown on Fig. 2.6: x has a positively skew distribution, $\log_e(x)$ has a negatively skew distribution, and $\log_e(x+\beta)$ has a symmetrical distribution.

The cumulative frequency distribution of a two-parameter lognormal variate will be plotted as a straight line on logarithmic-probability paper. If the variate is three-parameter lognormal the cumulative curve will show an excess of low values. The values of a three-parameter lognormal variate

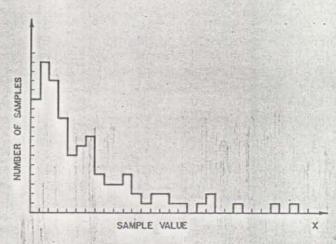


Fig. 2.5. Histogram of a positively skew distribution.

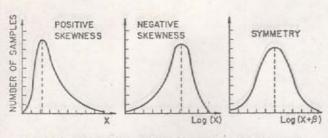


Fig. 2.6. Effect of the additive constant on the histogram.

are tabulated in Table 2.5 and represented graphically on Fig. 2.7 (curve C_1).

2.4.2 Parameter estimation and confidence intervals

The probability distribution of a three-parameter lognormal variate x is completely defined by: the additive constant β , the logarithmic variance of $(x + \beta)$, and the logarithmic mean of $(x + \beta)$. If we have n samples with values x_i (i = 1, 2, ..., n), we can estimate these three parameters,

Example of cumulative frequency distribution from a threeparameter lognormal variate

| Sample value upper limit (cm g) | Cumulative frequency (number of samples) | Cumulative frequency (%) | | |
|---------------------------------------|---|-----------------------------|--|--|
| 4 | 50 | 5 | | |
| 7 | 60 | 6 | | |
| 12 | 80 | 8 | | |
| 29 | 120 | 12 | | |
| 46 | 180 | 18 | | |
| 80 | 250 | 25 | | |
| 120 | 350 | 35 | | |
| 180 | 500 | 50 | | |
| 280 | 680 | 68 | | |
| 450 | 820 | 82 | | |
| 800 | 930 | 93 | | |
| 1 200 | 980 | 98 | | |
| 00 | 1 000 | 100 | | |

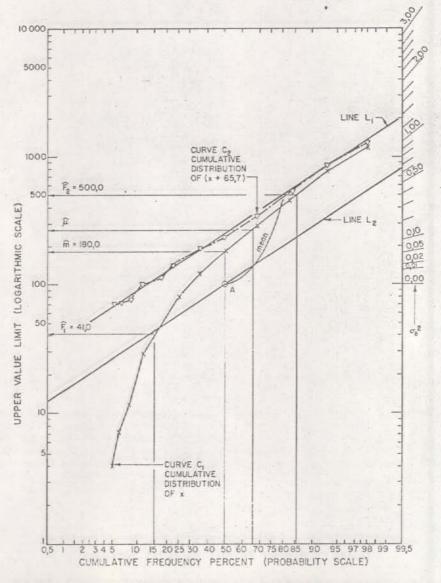


Fig. 2.7. Cumulative frequency distribution of a three-parameter lognormal variate.

from which we can calculate: An estimate $\hat{\mu}$ of the average value μ of the deposit; confidence limits for the average value of the deposit.

Estimation of the additive constant B

If the number of samples available is large enough, the cumulative frequency distribution is plotted as indicated on Fig. 2.7. Then β can be estimated using the following equation:

$$\beta = \frac{m^2 - f_1 f_1}{f_1 + f_2 - 2m},\tag{2.13}$$

where m is the sample value, corresponding to a 50% cumulative frequency (that is, the median of the observed distribution), and f_1 and f_2 are the sample values corresponding to p and 1 - p cumulative frequencies respectively. In theory, any value of p can be used but a value between 5% and 20% will give the best results. Consider Fig. 2.7. We read $\tilde{m} = 180.0$ cm g and for p = 15%, $f_1 = 41.0$ and $f_2 = 500.0$.

Hence $\hat{\beta} = 66$ cm g. Note that the value of $\hat{\beta}$ is often highly sensitive to the value of p chosen. It is therefore important to check graphically that the cumulative distribution of $x + \beta$

is lognormal (Fig. 2.7, curve C1).

If the number of samples n is small, it is not possible to estimate β graphically. Then one takes $\beta = 0$, or estimates β from the values this parameter takes in similar deposits. This last remark applies in particular to the case when a new gold mine is planned in the South African goldfields in a reef which has already been extensively mined in other areas. For examples of values of β , see Krige (1960).

Proof of Equation (2.13)

If $\log_{\epsilon}(x+\beta)$ is normally distributed, then because of the symmetry of the normal distribution about the mean, we can write:

$$\log_{\epsilon}(f_1 + \beta) + \log_{\epsilon}(f_2 + \beta) = 2\log_{\epsilon}(m + \beta),$$
 (2.14) which can be written:

$$(n_1 + \beta)^2 = (f_1 + \beta)(f_2 - \beta) \tag{2.15}$$

or
$$\beta = (m^2 - f_1 f_2)/(f_1 + f_2 - 2m)$$
. (2.16)

Estimation of the logarithmic mean and the geometric mean m From now on, we will assume that β is known. Let:

$$y_i = \log_e(x_i + \beta). \tag{2.17}$$

The natural logarithmic mean of $x - \beta$ is estimated from

$$\bar{y} = \frac{1}{n} \sum_{l=1}^{n} y_{l}. \tag{2.18}$$

The geometric mean m of $x + \beta$ is estimated by:

$$\hat{m} = \exp(\bar{y}). \tag{2.19}$$

Note that the geometric mean of a lognormal distribution is equal to the median of this distribution.

Estimation of the natural logarithmic variance V(y)

The variance estimator used for estimation of lognormal distributions is the maximum likelihood estimator, which differs from the unbiased estimator used for deriving normal distributions (2.5). Consider the normal variate:

$$y = \log_{e}(x + \beta). \tag{2.20}$$

The natural logarithmic variance σ_{ϵ}^2 of y is estimated by:

$$V(y) = \frac{1}{n} \sum_{i=1}^{n} (y_i - \bar{y})^2$$
 (2.21)

$$V(y) = \frac{1}{n} \sum_{i=1}^{n} y_i^2 - (\overline{y})^2$$
 (2.22)

Estimation of the average value of the deposit µ

The mean μ of a three-parameter lognormal variate x is From Fig. 2.8, given μ and the cumulative distribution line,

related to the geometric mean (or median) m and the natural logarithmic variance σ_c^2 by the formula:

$$\mu = m \left[\exp \left(\sigma_e^2 / 2 \right) \right] - \beta.$$
 (2.23)

When only estimates \hat{m} of m and V(y) of σ_e^2 are available, a maximum likelihood estimator for $(\mu + \beta)$ can be obtained using a table calculated by Sichel (1966, pp. 117-118, Table A). Part of this table is reproduced in Table 2.6. Given the number of samples n and the variance V = V(y), a factor $\gamma_n(V)$ can be read from these tables. The average sample value is then estimated by:

$$\hat{\mu} = \hat{m} \gamma_n (V) - \beta. \tag{2.24}$$

For n very large (> 1000) one can use:

$$\gamma_{\infty}(V) = \exp(V/2) \tag{2.25}$$

Example: For n = 10 and V = 1,40 we obtain from Table 26. $\gamma_n(V) = 1,936.$

Estimation of confidence limits of the mean µ

Tables have been developed by Sichel (1966, pp. 119-122) to calculate confidence limits of the mean of a lognormal distribution. These tables have been recalculated and expanded for $n \le 20$ by Wainstein (1975, pp. 228-238). Whenever possible the latter tables should be used as they are the more accurate, Sections of Sichel's and Wainstein's tables are reproduced in Tables 2.7 and 2.8.

We want to calculate the limit value μ_p such that the probability that μ is smaller than μ_p is p. From these tables and the corresponding values of p, n and V = V(y), a multiplying factor $\Psi_p(V;n)$ is obtained. We deduce:

$$\mu_p = (\hat{\mu} + \beta) \Psi_p (\nu; n) - \beta.$$
 (2.26)

For n very large (> 1000) one can use the following formula:

$$\Psi_p(V;n) = \exp(\sigma_t^2/2 + t_p \sigma_t),$$
 (2.27)

where
$$\sigma_t^4 = \frac{V}{n}(1 - \frac{V}{2})$$
 and t_p is obtained from Table 2.2.

For
$$p = 0.95$$
, $t_p = 1.645$ and for $p = 0.05$, $t_p = -1.645$.

Example: For p = 0.95, n = 10 and V = 1.40, we obtain from Table 2.8

$$\Psi_p(V;n) = 3,761.$$

2.4.3 Graphical estimation

Construction of a variance scale. Consider a lognormal variate x with median m = 100 and logarithmic variance $\sigma_e^2 = 1,00$. On logarithmic probability paper, the cumulative distribution of x is a straight line passing through the following two points (see § 2.3.3):

Point A: cumulative frequency = 0,50;

upper value limit = m = 100.

Point B: cumulative frequency = 0,97725;

upper value limit $= m \exp(2\sigma_e) = 738,9.$

This line is plotted on Fig. 2.8. The slope of the line is a function only of σ_e^2 . Leaving m constant and varying σ_e^2 , we obtain different lines through the point A, and can record the corresponding value of the variance on a scale as shown on the right-hand margin of Fig. 2.8.

Construction of the locus of the mean. The mean of x is given by (2.23). For m=100 and $\sigma_e^2=1,00$ we calculate:

$$\mu = m \exp(\sigma_{es}^2 2) = 164.9.$$

 $p_{\mu} = \operatorname{Prob}(x < \mu).$

| 7.7 | 3 |
|-----|---|
| ç | |
| = | |
| - | |
| | |
| | |
| | |

| | 1000 | 1,000 1,010 1,020 1,041 1,041 | 1,062 | 1,162 | 1,350 1,419 1,492 1,568 1,649 | 1,733 1,822 1,916 2,014 2,117 | 2,226 2,340 2,460 2,586 2,718 | 2,858 3,004 3,158 3,320 3,490 | 3,669 3,857 4,055 4,263 4,482 |
|--|------|--|---|-------|---|---|---|--|---|
| | 100 | 1,000 1,010 1,020 1,030 1,041 1,041 | 1,062 1,072 1,083 1,094 1,105 | 1,162 | 1,349 1,417 1,490 1,565 1,645 | 1,728 1,816 1,908 2,004 2,106 | 2,212 2,323 2,440 2,563 2,692 | 2,827 2,969 3,118 3,274 3,438 | 3,610 3,791 3,980 4,178 4,387 |
| | 50 | 1,000 1,010 1,020 1,030 1,041 1,051 | 1,062 1,072 1,083 1,094 1,105 | 1,161 | 1,348 1,416 1,487 1,562 1,641 | 1,723 1,810 1,900 1,995 2,095 | 2,199 2,308 2,422 2,542 2,668 | 2,800 2,937 3,082 3,233 3,391 | 3,557 3,730 3,912 4,102 4,301 |
| | 20 | 1,000 1,010 1,020 1,030 1,041 1,041 | 1,062 1,072 1,083 1,094 1,105 | 1,161 | 1,344 1,411 1,481 1,554 1,630 | 1,710 1,793 1,880 1,971 2,065 | 2,164 2,267 2,375 2,487 2,604 | 2,726 2,854 2,987 3,125 3,270 | 3,420 3,577 3,740 3,911 4,088 |
| | 18 | 1,000 1,010 1,020 1,030 1,041 1,051 | 1,062 1,072 1,094 1,104 | 1,160 | 1,344 1,410 1,480 1,552 1,628 | 1,708 1,790 1,876 1,966 2,060 | 2,158 2,260 2,367 2,478 2,594 | 2,714 2,840 2,971 3,108 3,250 | 3,398 3,552 3,713 3,880 4,054 |
| 01,1 = (07 | 91 | 1,000 1,010 1,020 1,030 1,041 1,051 | 1,062 1,072 1,083 1,094 1,104 | 1,160 | 1,343 1,409 1,478 1,551 1,626 | 1,705 1,787 1,872 1,961 2,054 | 2,151 2,252 2,357 2,467 2,581 | 2,700 2,824 2,952 3,086 3,226 | 3,371 3,522 3,680 3,843 4,013 |
| Example: $\gamma_a(0,20) = 1,104$. | 14 | 1,000 1,010 1,020 1,030 1,041 | 1,062 1,072 1,083 1,094 1,104 | 1,160 | 1,342 1,408 1,477 1,549 1,623 | 1,701 1,782 1,867 1,955 2,047 | 2,142 2,242 2,345 2,453 2,565 | 2,682 2,803 2,929 3,060 3,197 | 3,339 3,486 3,639 3,799 3,964 |
| | 12 | 1,000 1,010 1,020 1,030 1,041 1,051 | 1,062 1,072 1,083 1,094 1,104 | 1,160 | 1,341 1,406 1,475 1,546 1,620 | 1,697 1,777 1,860 1,947 2,037 | 2,131 2,229 2,330 2,436 2,545 | 2,659 2,778 2,900 3,028 3,160 | 3,298 3,441 3,589 3,743 3,902 |
| opulation. | 01 | 1,000 1,010 1,020 1,030 1,041 | 1,061 1,072 1,083 1,093 1,104 | 1,160 | 1,339 1,404 1,472 1,542 1,615 | 1,691 1,770 1,851 1,936 2,025 | 2,116 2,212 2,310 2,413 7,519 | 2,630 2,744 2,863 2,986 3,113 | 3,245 3,382 3,524 3,671 3,824 |
| gnormal p | 6 | 1,000 1,010 1,020 1,030 1,041 1,051 | 1,061 1,072 1,083 1,093 1,104 | 1,159 | 1,338 1,403 1,470 1,540 1,612 | 1,687 1,765 1,846 1,930 2,017 | 2,107 2,201 2,298 2,399 2,503 | 2,611 2,723 2,839 2,959 3,084 | 3,213 3,346 3,484 3,627 3,775 |
| mean of lo | 00 | 1,000 1,010 1,020 1,030 1,041 1,041 | 1,061 1,072 1,082 1,093 1,104 | 1,159 | 1,337 1,401 1,468 1,537 1,608 | 1,682 1,759 1,839 1,922 2,007 | 2,096 2,188 2,283 2,382 2,484 | 2,589 2,698 2,811 2,928 3,049 | 3,174 3,304 3,437 3,576 3,718 |
| Table 2.0 man of lognormal population. | 7 | 1,000 | 1,061 1,072 1,082 1,093 1,104 | 1,159 | 1,336 1,399 1,465 1,533 1,604 | 1,677 1,752 1,831 1,912 1,996 | 2,082 2,172 2,265 2,361 2,460 | 2,563 2,669 2,778 2,891 3,008 | 3,128 3,253 3,382 3,514 3,651 |
| V) for estin | 9 | 1,000 | 1,061 1,072 1,082 1,093 1,104 | 1,158 | 1,334 1,397 1,462 1,529 1,598 | 1,670 1,744 1,820 1,900 | 2,066 2,153 2,243 2,336 2,431 | 2,530 2,632 2,737 2,846 2,957 | 3,073 3,191 3,314 3,440 3,570 |
| Factor $\gamma_n(V)$ for esti | S | 1,000 1,010 1,020 1,030 1,040 | 1,061 1,072 1,082 1,093 | 1,158 | 1,332 1,393 1,457 1,523 1,591 | 1,661 1,733 1,807 1,884 1,963 | 2,044 2,128 2,214 2,303 2,395 | 2,489 2,586 2,686 2,788 2,788 | 3,003 3,114 3,229 3,347 3,469 |
| 11 | 4 | 1,000 | 1,061 1,071 1,082 1,092 1,103 | 1,157 | 1,328 1,389 1,451 1,515 | 1,648 1,718 1,789 1,862 | 2,015 2,095 2,177 2,260 2,347 | 2,435 2,526 2,618 2,714 2,812 | 2,912 3,015 3,120 3,228 3,339 |
| | т. | 1,000 | 1,061 1,071 1,082 1,092 1,102 | 1,156 | 1,323 1,382 1,442 1,503 1,566 | 1,630 1,696 1,764 1,832 1,903 | 1,975 2,049 2,124 2,201 2,280 | 2,360 2,442 2,526 2,612 2,612 2,699 | 2,789 2,880 2,973 3,068 3,166 |
| | 2 | 1,000 | 1,061 1,081 1,081 1,091 1,091 | 1,154 | 1,315 1,371 1,427 1,485 1,543 | 1,602 1,662 1,724 1,786 | 1,912 1,977 2,043 2,110 2,178 | 2,247 2,317 2,388 2,460 2,533 | 2,687 2,682 2,759 2,836 2,914 |
| | | | | | | | | | |
| | 1 2 | 0,00 | 0,12 0,14 0,16 0,16 0,18 | 0.3 | 0,0 | : ::: | 6,1,1,1,1,0,2 | 22.2 | 0,11,11,15, 0,15, |

CLASSICAL STATISTICS, NORMAL AND LOGNORMAL THEORY

(From H. S. Sichel, Journal of the South African Institute of Mining and Metallurgy, March 1966)

Factor $\Psi_{0,93}$ (V:n) for estimation of one-sided upper 95% confidence limits of the mean of a lognormal population.

| 0,00 0,02 0,04 0,06 | 1,000 1,241 1,362 1,466 1,561 | 1,000 1,117 1,171 | 1,000 1,084 | 1,000 | 1,000 | 1,000 | Marie and |
|------------------------------|---|-------------------------|----------------|-------|-------|----------------|----------------|
| 0,02 | 1,241 1,362 1,466 1,561 | 1,117 | | | | 1.11111 | 1,000 |
| 0,04 | 1,362 1,466 1,561 | | | 1,067 | 1,038 | 1,026 | 1,007 |
| | 1,466 1,561 | | 1,122 | 1,099 | 1,055 | 1,037 | 1,011 |
| 11110 | 1,561 | 1,216 | 1,154 | 1,124 | 1,069 | 1,046 | 1,013 |
| 0,08 | | 1,256 | 1,181 | 1,146 | 1,080 | 1,053 | 1,015 |
| 0,10 | 1,652 | 1,293 | 1,207 | 1,166 | 1,091 | 1,060 | 1,017 |
| The second | 1,002 | 1,223 | 1,207 | 1,100 | 1,051 | 1,000 | 1,017 |
| 0,12 | 1,740 | 1,327 | 1,230 | 1,184 | 1,100 | 1,066 | 1,019 |
| 0,14 | 1,827 | 1,361 | 1,253 | 1,202 | 1,109 | 1,072 | 1,020 |
| 0,16 | 1,914 | 1,393 | 1,274 | 1,219 | 1,118 | 1,078 | 1,022 |
| 0,18 | 1,999 | 1,425 | 1,295 | 1,236 | 1,126 | 1,084 | 1,023 |
| 0,20 | 2,087 | 1,455 | 1,316 | 1,252 | 1,135 | 1,089 | 1,025 |
| 0,30 | 2,532 | 1,606 | 1,415 | 1,328 | 1,172 | 1,113 | 1,031 |
| 0,40 | 3,019 | 1,756 | 1,509 | 1,399 | 1,207 | 1,135 | 1,037 |
| 0,50 | 3,563 | 1,910 | 1,603 | 1,470 | 1,240 | 1,156 | 1,042 |
| 0,60 | 4,176 | 2,070 | 1,682 | 1,541 | 1,273 | 1,175 | 1,047 |
| 0,70 | 4,870 | 2,237 | 1,798 | 1,614 | 1,306 | 1,196 | 1,052 |
| 0,80 | 5,663 | 2,415 | 1,901 | 1,688 | 1,338 | 1,215 | 1,057 |
| 0,90 | 6,570 | 2,604 | 2,006 | 1,763 | 1,371 | 1,235 | 1,062 |
| 1,00 | 7,605 | 2,805 | 2,117 | 1,842 | 1,404 | 1,254 | 1,067 |
| 1,10 | 8,795 | 3,019 | 2,233 | 1,924 | 1,437 | 1,274 | 1,071 |
| 1,20 | 10,155 | 3,250 | 2,355 | 2,008 | 1,471 | 1,294 | 1,076 |
| 1,30 | 11,718 | 3,497 | 2,483 | 2,096 | 1,506 | 1,314 | 1,080 |
| 1,40 | 13,513 | 3,761 | 2,617 | 2,187 | 1,540 | 1,334 | 1,085 |
| 1,50 | 15,569 | 4,045 | 2,758 | 2,282 | 1,576 | 1,354 | 1,089 |
| 1,60 | 17,928 | 4,351 | 2,907 | 2,380 | 1,613 | 1,374 | 1,094 |
| 1,70 | 20,639 | 4,680 | 3,064 | 2,484 | 1,650 | 1,395 | 1,098 |
| 1,80 | 23,749 | 5,034 | 3,229 | 2,592 | 1,688 | 1,416 | 1,103 |
| 1,90 | 27,318 | 5,414 | 3,403 | 2,704 | 1,728 | 1,438 | 1,107 |
| 2,00 | 31,398 | 5,825 | 3,588 | 2,822 | 1,767 | 1,459 | 1,112 |
| 2,10 | 36,079 | 6,268 | 3,783 | 2,945 | 1,808 | 1,481 | 1,116 |
| 2,20 | 41,444 | 6,745 | 3,989 | 3,074 | 1,850 | 1,504 | 1,121 |
| 2,30 | 47,586 | 7,260 | 4,208 | 3,209 | 1,893 | 1,526 | 1,125 |
| 2,40 | 54,611 | 7,815 | 4,438 | 3,351 | 1,937 | 1,549 | 1,130 |
| 2,50 | 62,661 | 8,415 | 4,683 | 3,498 | 1,982 | 1,572 | 1,134 |
| 2,60 | 71,861 | 9,061 | 4,941 | 3,670 | 2,029 | 1,596 | 1,139 |
| 2,70 | 82,366 | 9,759 | 5,214 | 3,816 | 2,076 | 1,620 | 1,144 |
| 2,80 | 94,377 | 10,512 | 5,504 | 3,986 | 2,076 | | |
| 2,90 | 108,115 | 11,326 | 5,811 | 4,164 | 2,175 | 1,645 | 1,148 |
| 3,00 | 123,750 | 12,206 | 6,137 | 4,104 | 2,175 | 1,670 1,695 | 1,153 1,158 |

Example: $\Psi_{0,95}$ (1,60;15) = 2,907

(From H. S. Sichel, March 1966, and B. M. Wainstein, April 1975, Journal of the South African Institute of Mining and Metallurgy)

we read $p_{\mu}=69.5\%$. If we vary σ_{ℓ}^{2} , leaving m=100 constant, the point M with co-ordinates (p_{μ}, μ) will describe a curve which is the locus of the mean.

Graphical estimation of mean and variance. One must use logarithmic probability paper on which the variance scale and the locus of the mean have been drawn. Consider the variate presented in Table 2.5. The following procedure must be used (see Fig. 2.7).

- Draw the observed cumulative distribution of x (curve C₁).
- Estimate β (§ 2.4.2). Let $\hat{\beta}$ be the estimate of β .
- Draw the cumulative distribution of $x + \hat{\beta}$ (curve C_2).
- Fit a straight line L1 to the curve C2.

- Draw a line L₂ parallel to L₁ through the reference point A
 which was used to draw the variance scale and mean locus.
- Read the logarithmic variance on the variance scale: $\hat{\sigma}_{e}^{2} = 0.65$.
- Project the intersection of L₁ with the locus of the mean onto line L₁ and determine the mean:

$$\hat{\mu} + \hat{\beta} = 350$$
 $\hat{\mu} = 350 - 66 = 284$.

Note that exact confidence limits cannot be obtained from these estimates, but in practice (2.26) can be used, substituting $\hat{\sigma}_e^z$ for V.

Table 2.8

Factor $\Psi_{0,05}(V;n)$ for estimation of one-sided lower 95% confidence limits of the mean of a lognormal population.

| v | 5 | 10 | 15 | 20 | 50 | 100 | 1000 |
|--------|--------|--------|--------|------------------|------------------|---------------|--------|
| 0,00 | 1,0000 | 1,0000 | 1,0000 | 1,0000 | 1,0000 | 1,0000 | 1,0000 |
| 0,02 | 0,8978 | 0,9333 | 0,9458 | 0,9540 | 0,9697 | 0,9782 . | 0,9927 |
| 0,04 | 0,8589 | 0,9071 | 0,9246 | 0,9344 | 0,9573 | 0,9692 | 0,9895 |
| 0.06 | 0,8302 | 0,8874 | 0,9079 | 0,9200 | 0,9478 | 0,9622 | 0,9872 |
| 0,08 | 0,8070 | 0,8708 | 0,8943 | 0,9077 | 0,9398 | 0,9564 | 0,9852 |
| 0,10 | 0,7870 | 0,8563 | 0,8821 | 0,8972 | 0,9328 | 0,9512 | 0,9833 |
| 0,12 | 0,7693 | 0,8439 | 0,8716 | 0,8878 | 0,9264 | 0,9464 | 0,9817 |
| 0,14 | 0,7535 | 0,8323 | 0,8617 | 0,8790 | 0,9204 | 0,9420 | 0,9801 |
| 0,16 | 0,7389 | 0,8216 | 0,8527 | 0,8709 | | 0,9420 | |
| 0,18 | 0,7255 | 0,8116 | 0,8442 | | 0,9149 | 0,9380 | 0,9787 |
| 0,20 | 0,7129 | 0,8023 | 0,8360 | 0,8632 0,8558 | 0,9097 0,9048 | 0,9341 0,9304 | 0,9773 |
| 0,30 | 0,6605 | 0.7619 | 0.8008 | 0.0242 | 0.0000 | | |
| | | 0,7618 | 0,8008 | 0,8243 | 0,8828 | 0,9139 | 0,9701 |
| 0,40 | 0,6187 | 0,7284 | 0,7717 | 0,7981 | 0,8639 | 0,8996 | 0,9648 |
| 0,50 | 0,5838 | 0,6995 | 0,7462 | 0,7744 | 0,8470 | 0,8867 | 0,9600 |
| 0,60 | 0,5538 | 0,6739 | 0,7270 | 0,7534 | 0,8313 | 0,8741 | 0,9554 |
| 0,70 | 0,5277 | 0,6508 | 0,7020 | 0,7338 | 0,8168 | 0,8632 | 0,9511 |
| 0,80 | 0,5044 | 0,6297 | 0,6825 | 0,7156 | 0,8030 | 0,8525 | 0,9470 |
| 0,90 | 0,4836 | 0,6103 | 0,6646 | 0,6987 | 0,7899 | 0,8421 | 0,9429 |
| 00,1 | 0,4650 | 0,5923 | 0,6476 | 0,6826 | 0,7774 | 0,8322 | 0,9389 |
| 1,10 | 0,4481 | 0,5756 | 0,6317 | 0,6674 | 0,7654 | 0,8226 | 0,9351 |
| 1,20 | 0,4328 | 0,5599 | 0,6165 | 0,6530 | 0,7538 | 0,8133 | 0,9313 |
| 1,30 | 0,4189 | 0,5452 | 0,6023 | 0,6393 | 0,7426 | 0,8042 | 0,9276 |
| 1,40 | 0,4062 | 0,5315 | 0,5888 | 0,6262 | 0,7318 | 0,7954 | 0,9240 |
| 1,50 | 0,3946 | 0,5186 | 0,5760 | 0,6137 | 0,7214 | 0,7868 | 0,9203 |
| 1,60 | 0,3840 | 0,5065 | 0,5637 | 0,6018 | 0,7112 | 0,7784 | 0,9168 |
| 1,70 | 0,3743 | 0,4950 | 0,5521 | 0,5904 | 0,7014 | 0,7704 | |
| 1,80 | 0,3655 | 0,4842 | 0,5410 | | | 0,7702 | 0,9133 |
| 1,90 | 0,3574 | 0,4740 | | 0,5794 | 0,6918 | 0,7622 | 0,9098 |
| 2,00 | | | 0,5305 | 0,5688 | 0,6825 | 0,7544 | 0,9064 |
| 2,00 | 0,3501 | 0,4644 | 0,5203 | 0,5587 | 0,6734 | 0,7466 | 0,9030 |
| 2,10 | 0,3433 | 0,4552 | 0,5106 | 0,5489 | 0,6646 | 0,7391 | 0,8996 |
| 2,20 | 0,3372 | 0,4466 | 0,5014 | 0,5395 | 0,6560 | 0,7317 | 0,8962 |
| 2,30 | 0,3316 | 0,4385 | 0,4925 | 0,5304 | 0,6476 | 0,7245 | 0,8929 |
| 2,40 | 0,3266 | 0,4308 | 0,4840 | 0,5217 | 0,6394 | 0,7173 | 0,8896 |
| 2,50 - | 0,3220 | 0,4234 | 0,4759 | 0,5133 | 0,6314 | 0,7104 | 0,8864 |
| 2,60 | 0,3179 | 0,4166 | 0,4681 | 0,5044 | 0,6236 | 0,7035 | 0,8831 |
| 2,70 | 0,3142 | 0,4100 | 0,4606 | 0,4974 | 0,6160 | 0,6967 | 0,8799 |
| 2,80 | 0,3110 | 0,4039 | 0,4535 | 0,4899 | 0,6085 | 0,6901 | 0,8767 |
| 2,90 | 0,3081 | 0,3981 | 0,4467 | 0,4826 | 0,6012 | 0,6836 | 0,8736 |
| 3,00 | 0,3055 | 0,3926 | 0,4401 | 0,4756 | 0,5941 | 0,6772 | 0,8704 |

Example: $\Psi_{0.05}$ (1,60:15) = 0,5637

(From H. S. Sichel, March 1966, and B. M. Wainstein, April 1975, Journal of the South African Institute of Mining and Metallurgy)

2.4.4 Example

Consider the same data as in §2.3.4 and assume that x has a two-parameter lognormal distribution ($\beta=0$). Estimate the mean μ of the ore body and give the central 90% confidence limits for this mean.

Using Table 2.9, we calculate:

$$\bar{y} = 4,581/8 = 0,573,$$

 $\hat{m} = \exp(\bar{y}) = 1,774,$
 $V(y) = 2,983/8 - (0,573)^2 = 0,0445.$

Using Table 2.6 we obtain by interpolation between the points corresponding to V = 0.04 and V = 0.06 for n = 8: $\gamma_n(V) = 1.022$.

Hence the estimate of the mean of the ore body:

$$\hat{\mu} = 1,774 \times 1,022 = 1,830.$$

Table 2.9
Calculation of logarithmic mean and variance

| i | XI | $y_i = \log_e(x_i)$ | yi |
|-------|-----|----------------------|------------------------|
| 1 | 1,2 | 0,182 | 0,033 |
| 2 | 2,0 | 0,693 | 0,481 |
| 3 | 1,6 | 0,470 | 0,221 |
| 4 | 1,7 | 0,531 | 0,282 |
| 5 | 2,5 | 0,916 | 0,840 |
| 6 | 1,9 | 0,642 | 0,412 |
| 7 | 1,5 | 0,405 | 0,164 |
| 8 | 2,1 | 0,742 | 0,550 |
| n = 8 | | $\Sigma y_l = 4,581$ | $\Sigma y_i^* = 2,983$ |

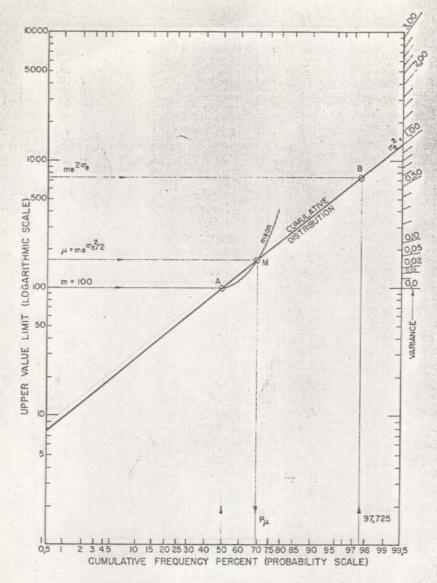


Fig. 2.8. Construction of variance scale and mean locus.

The central 90% confidence limits are:

upper limit: $\hat{\mu}_{0,85} = \hat{\mu} \Psi_{0.95} (V;n)$

lower limit: $\hat{\mu}_{0,05} = \hat{\mu} \Psi_{0,05} (V;n)$ Using Tables 2.7 and 2.8, we obtain by linear interpolation between the points corresponding to V = 0.04 and V = 0.06,

n = 5 and n = 10:

 $\Psi_{0,us}(V;n) = 1,263$ $\Psi_{0,us}(V;n) = 0,883$

Hence:

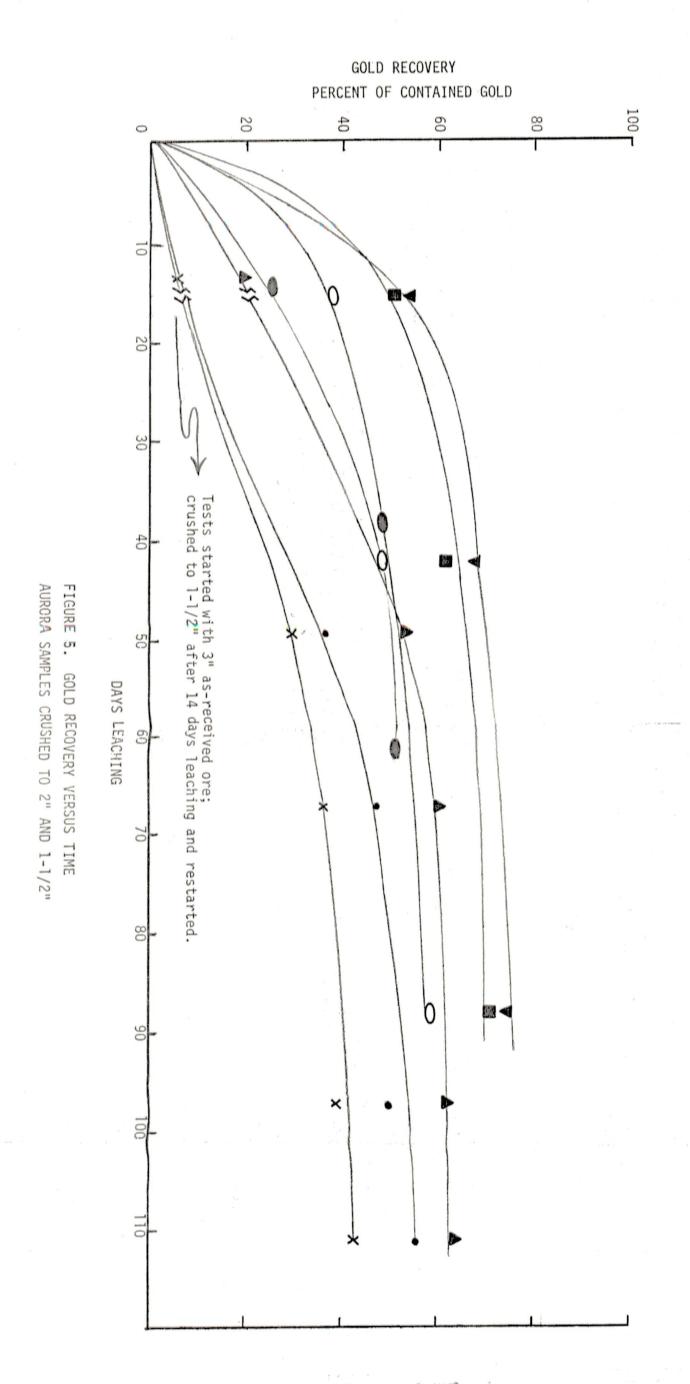
 $\mu_{0,95} = 1,263 \times 1,830 = 2,311$

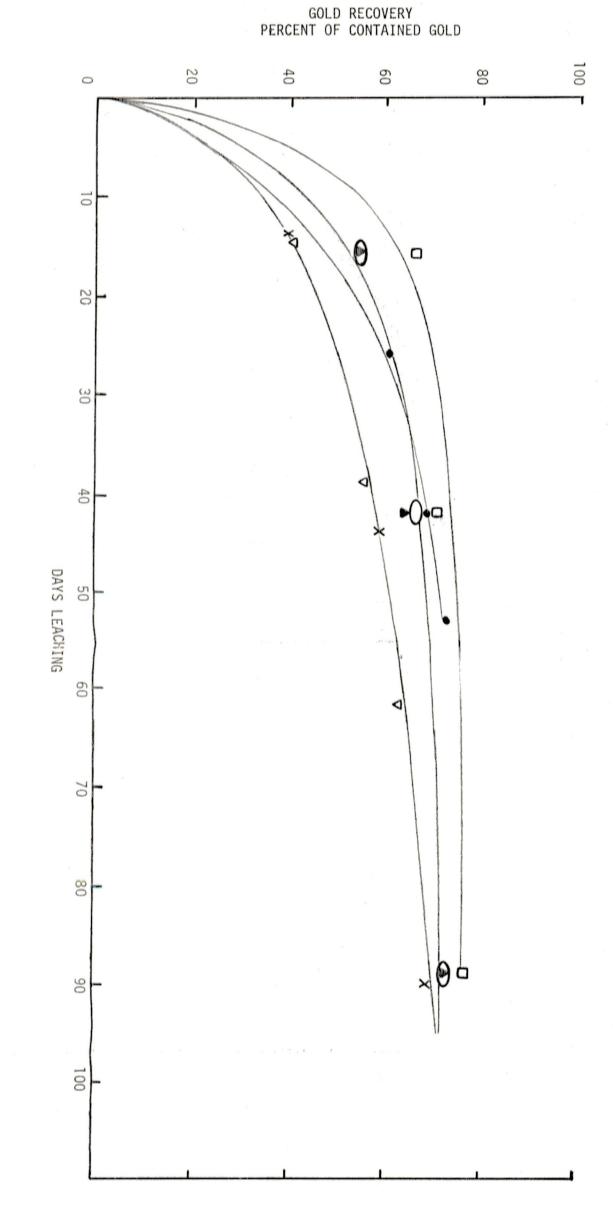
 $\mu_{0*03} = 0,883 \times 1,830 = 1,615$

Compare these limits with those obtained in § 2.3.4.

FIGURE 1. SAMPLE DESCRIPTIONS AND HEAD ASSAYS

| MILLER-KAPPES SAMPLE NUMBER | LOCATION | DESCRIPTION | 20. d | HEAD ASSAY oz Au/Ton |
|--------------------------------|---|---|----------|-------------------------|
| 1701 A | Western edge stope, Humboldt East vein, bottom of gloryhole | Hand-picked rocks; a few fines, but mostly in 2-inch to 4-inch range | | |
| 1701 В | Humboldt East vein, westernmost stope, upper East rib | Same description as 1701 A | | .105 |
| 1701 C | Same as 1701 B | Same description as 1701 A | | .105 |
| 1701 D | Humboldt East vein, central zone, vertical shaft dump | Taken off of surface of white dump surrounding vertical shaft. Some rocks up to 4-inches, but mostly 1/2-inch to 2-inches with some fines | | |
| 1759 A | N28E and 460 feet from southern common corner Curry and East Humboldt claim | High grade vein matter at surface, 6 to 8 feet thickness of footwall at main vein | | .199 |
| 1759 В | N55W from sample and 50 feet to drillhole PH #1 | Modest grade quartz vein, surface sample collected from 3 to 9 feet of hanging wall side of vein | | .011 |
| 1759 C | Stope at Kurt Guide Station, 1400 + 147 NW | Good grade quartz matter. Sample taken across 20 feet of footwall of vein, excluding immediate 5 feet of footwall quartz. Collected from 20 to 40 feet below surface. | | .079 |
| 1759 D | Humboldt East dump. Old collapsed adit on northerly end and west side of main vein. | Oxidized material hand-picked off dump | | .081 |
| 1759 E | Same as 1759 D | Pyritic/unoxidized ore, hand-picked off dump | | .089 |
| 1902 | East Humboldt adit | Collected off of the dump of "long drift" on the East side of vein. Sample from approximately 50 to 75 feet below the surface. | | .412 |





• 1737 • 1770 • 1770 • 1772 • 1774 × 1795 • 1960

FIGURE 6. GOLD RECOVERY VERSUS TIME AURORA SAMPLES CRUSHED TO 5/8-INCHES

| TEST | CALC.HEAD ASSAY | SIZE FRACTION | | | | | | |
|-------------------|-----------------|---------------------|---------------------|---------------------|--------------------|--------------------|--------------------|----------------|
| NO. | oz Au/Ton | + 1" | -1" + 1/2" | -1/2" + 3M | -3M + 10M | -10M +65M | -65M | TOTAL |
| 1702 | .366 | 13,780 .260/.224 | 2,820 .184/.168 | 1,210 .128/.120 | 1,660 .110/.100 | | | 19,470 .213 |
| 1703 | .119 | 7,450 .054/.040 | 5,180 .076/.068 | 1,900 .056/.046 | 1,360 .038/.039 | 580 .040/.036 | 300 .078/.006 | 16,770 .054 |
| 1704 | .054 | 11,460 .025/.025 | 4,760 .016/.013 | 2,080 .014/.017 | 2,010 .017/.016 | 970 .008/.013 | 440 .026/.014 | 21,720 .020 |
| 1736 | .109 | | | | | | | 3,150 .023 |
| 1737 ² | .116 | | | | MR 000 000 | | | 3,360 .031 |
| 1769 | .143 | 12,180 .042/.054 | 4,640 .038/.045 | 1,920 .035/.036 | 1,840 .031/.033 | 1,200 .019/.024 | 630 .023/.024 | 22,410 .042 |
| 1771 | .011 | 12,340 .004/.004 | 3,990 .002/.002 | 1,480 .002/.003 | 1,250 .002/.001 | 870 .002/.003 | | 19,930 .003 |
| 1772 | .009 | | 2,680 .000/.007 | 9,250 .000/.004 | 5,050 .000/.002 | 2,060 .000/.001 | 880 .000/.007 | 19,920 .002 |
| 1773 | .098 | 11,420 .048/.042 | 4,970 .042/.046 | 1,260 .030/.025 | 1,420 .026/.024 | 1,090 .034/.043 | | 20,160 .042 |
| 1774 | .085 | | 2,120 .032/.023 | 9,390 .026/.032 | 5,730 .018/.015 | 2,060 .014/.011 | 740 .016/.014 | 20,040 |
| .795 | .083 | | 4,040 .050/.036 | 11,240 .028/.024 | 5,270 .018/.019 | 1,860 .018/.010 | 1,080 .012/.010 | 23,490 .026 |
| .959 | .395 | 13,720 .232/.206 | 3,610 .180/.134 | 1,350 .156/.137 | 1,190 .106/.108 | - | | 19,870 .196 |
| .960 | .428 | NO 400 100 | 12,370 .190/.176 | 4,430 .148/.125 | 2,800 .138/.113 | 1,060 .086/.078 | 450 .090/.083 | 21,110 .158 |

¹⁻Weight in grams; duplicate fire assays (ounces Au per short ton).